Optimal experimental design for Bayesian inverse problems governed by PDE models with uncertainty with application to subsurface flow and tsunami equations

Karina Koval

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Motivating examples

Introduction and background

OED for Gaussian regression OED for Bayesian inverse problems

OED under model uncertainty

Mathematical formulation of OED Computational challenges Numerical results – subsurface flow

OED for tsunami source reconstruction Mathematical formulation Numerical results

Summary

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Motivation

$$oldsymbol{d} = \mathcal{F}(oldsymbol{m}) + oldsymbol{\eta}$$

- \blacktriangleright Models for real-world phenomena involve unknown parameters, m
- Accurate estimation of parameters relies on informative data, d (inverse problem)
- Data collection limited due to cost or physical constraints
- Optimal experimental design (OED): Design of experimental conditions for parameter inference problems governed by PDE – where and what to measure/observe?

Contaminants in groundwater

Inverse problem:

Given concentration readings, infer source of contamination



Graphic from sciencefriday.com

Contaminants in groundwater

Inverse problem:

Given concentration readings, infer source of contamination

OED problem:

Where to drill wells to optimally infer initial source in event of contamination



Graphic from sciencefriday.com

Contaminants in groundwater - governing model

$$u_t - \kappa \Delta u + \boldsymbol{v} \cdot \nabla u = 0$$
$$u(\cdot, T_i) = \boldsymbol{m}$$
$$+ BCs$$

- $u(\boldsymbol{x},t)$: concentration
- $\kappa > 0$: diffusion coefficient
- v(x,t): advection velocity field
- ▶ *m*: unknown initial concentration



Earthquake-generated tsunamis

- Tsunamis generated by earthquakes beneath ocean floor
- Earthquake ~> ocean floor deformation ~> tsunami waves
- Tsunami warning relies on knowledge of bathymetry change
- Cannot measure this, but can measure water depth



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Inverse problem: Given water level measurements, reconstruct ocean floor deformation

OED problem: Place sensors for optimal tsunami source reconstruction and accurate tsunami forecasting



Earthquake-generated tsunamis - governing equations

$$\begin{bmatrix} h \\ u \\ v \end{bmatrix}_t + \begin{bmatrix} u \\ \frac{u^2}{h} + \frac{1}{2}gh^2 \\ \frac{uv}{h} \end{bmatrix}_x + \begin{bmatrix} v \\ \frac{uv}{h} \\ \frac{v^2}{h} + \frac{1}{2}gh^2 \end{bmatrix}_y = - \begin{bmatrix} 0 \\ ghB_x \\ ghB_y \end{bmatrix}$$

•
$$h(x, y, t)$$
: water depth

• u(x, y, t) and v(x, y, t): fluid momentum

- B(x,y): bathymetry
- Models wave propagation due to bathymetry change



(Simulations performed in GeoClaw)

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Gaussian process regression

Measure $m(x_i)$ directly at points x_i



$$d_i = m(x_i) + \eta_i$$

• Goal: Determine distribution for m(x)given noisy data $d \in \mathbb{R}^s$

► Gaussian process: probabilistic approach to regression problems

Uses (noisy) data d to update prior knowledge about m ~> posterior distribution

$$d_i = m(x_i) + \eta_i$$

Goal: determine $m^* = [m(x_1^*), \dots, m(x_n^*)]$ for $x_i^* \in D^*$ **Given:** data $d = [d_1, \dots, d_s]$ at x_1, \dots, x_s for $x_i \in D_d$

$$d_i = m(x_i) + \eta_i$$

Goal: determine $m^* = [m(x_1^*), \dots, m(x_n^*)]$ for $x_i^* \in D^*$ **Given:** data $d = [d_1, \dots, d_s]$ at x_1, \dots, x_s for $x_i \in D_d$

Assume $m \sim \mathcal{N}(0, \mathcal{C}_{pr})$ (prior) and $\eta_i \sim \mathcal{N}(0, \sigma_n^2)$

• C_{pr} defined through *covariance function* c(x, y)



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Assume m ~ N(0, C_{pr}) (prior) and η_i ~ N(0, σ_n²)
 C_{pr} defined through covariance function c(x, y)



$$d_i = m(x_i) + \eta_i$$

Goal: determine $\boldsymbol{m}^* = [\boldsymbol{m}(x_1^*), \dots, \boldsymbol{m}(x_n^*)]$ for $x_i^* \in D^*$ **Given:** data $\boldsymbol{d} = [d_1, \dots, d_s]$ at x_1, \dots, x_s for $x_i \in D_{\boldsymbol{d}}$ Assume $\boldsymbol{m} \sim \mathcal{N}(0, \mathcal{C}_{pr})$ (prior) and $\eta_i \sim \mathcal{N}(0, \sigma_n^2)$

→ Joint multivariate distribution:

$$\begin{bmatrix} \boldsymbol{d} \\ \boldsymbol{m}^* \end{bmatrix} \sim \mathcal{N} \begin{pmatrix} \mathbf{0}, \begin{bmatrix} \mathbf{C}_{\mathsf{pr}}(D_{\boldsymbol{d}}, D_{\boldsymbol{d}}) + \sigma_n^2 \mathbf{I} & \mathbf{C}_{\mathsf{pr}}(D_{\boldsymbol{d}}, D^*) \\ \mathbf{C}_{\mathsf{pr}}(D^*, D_{\boldsymbol{d}}) & \mathbf{C}_{\mathsf{pr}}(D^*, D^*) \end{bmatrix} \end{pmatrix},$$

where, e.g., $\mathbf{C}_{pr}(D^*, D_d) \in \mathbb{R}^{n \times s}$ with $[\mathbf{C}_{pr}(D^*, D_d)]_{ij} = c^{SE}(x_i^*, x_j)$

$$d_i = m(x_i) + \eta_i$$

Goal: determine $m^* = [m(x_1^*), \dots, m(x_n^*)]$ for $x_i^* \in D^*$ **Given:** data $d = [d_1, \dots, d_s]$ at x_1, \dots, x_s for $x_i \in D_d$ Assume $m \sim \mathcal{N}(0, \mathcal{C}_{pr})$ (prior) and $\eta_i \sim \mathcal{N}(0, \sigma_n^2)$

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Bayesian inference \rightsquigarrow posterior $m^*|d \sim \mathcal{N}(m_{\mathsf{post}}, \mathbf{C}_{\mathsf{post}})$:

$$\begin{split} \boldsymbol{m}_{\mathsf{post}} &= \mathbf{C}_{\mathsf{pr}}(D^*, D_d) \left[\mathbf{C}_{\mathsf{pr}}(D_d, D_d) + \sigma_n^2 \mathbf{I} \right]^{-1} d \\ \mathbf{C}_{\mathsf{post}} &= \mathbf{C}_{\mathsf{pr}}(D^*, D^*) - \mathbf{C}_{\mathsf{pr}}(D^*, D_d) \left[\mathbf{C}_{\mathsf{pr}}(D_d, D_d) + \sigma_n^2 \mathbf{I} \right]^{-1} \mathbf{C}_{\mathsf{pr}}(D_d, D^*) \end{split}$$

Gaussian regression - samples and variance





 $\begin{array}{l} \textbf{Prior:} \\ \boldsymbol{m^{*}} \sim \mathcal{N}\left(\boldsymbol{0}, \boldsymbol{\mathrm{C}_{\mathrm{pr}}}(D^{*}, D^{*}) \right) \end{array}$

Posterior: $m{m}^* | m{d} \sim \mathcal{N}(m{m}_{\mathsf{post}}, \mathbf{C}_{\mathsf{post}})$

OED for Gaussian regression



Can observe d at a limited number of locations of our choice
 How to choose these locations to optimally infer m*?

OED for Gaussian regression



- \blacktriangleright Can observe d at a limited number of locations of our choice
- How to choose these locations to optimally infer m^* ?

Requires:

- $1. \ \ \text{Incorporation of design}$
- 2. Description of "optimal" design
- 3. Incorporation of cost constraints

1. Definition and incorporation of design

Design definition is problem specific

For 1D Gaussian regression:



▶ Assume grid of s possible measurement locations, $x_i \in [a, b]$

• Assign binary weight w_i to measurement at location x_i

 $w_i = \begin{cases} 1 \implies \text{ use measurement at } x_i \\ 0 \implies \text{ ignore measurement at } x_i \end{cases}$

 $\blacktriangleright \boldsymbol{w} = [w_1, \ldots, w_s]$

1. Definition and incorporation of design

Design-dependent model:

$$oldsymbol{d}(oldsymbol{w}) = \mathbf{W}\left(oldsymbol{m}+oldsymbol{\eta}
ight)$$

•
$$\mathbf{W} := \mathbf{W}(\boldsymbol{w}) \in \mathbb{R}^{k(\boldsymbol{w}) \times s}$$
:

$$\begin{bmatrix} \boldsymbol{d}(\boldsymbol{w}) \\ \boldsymbol{m}^*(\boldsymbol{w}) \end{bmatrix} \sim \mathcal{N} \left(\mathbf{0}, \begin{bmatrix} \mathbf{W} \left(\mathbf{C}_{\mathsf{pr}}(D_{\boldsymbol{d}}, D_{\boldsymbol{d}}) + \sigma_n^2 \mathbf{I} \right) \mathbf{W}^T & \mathbf{W} \mathbf{C}_{\mathsf{pr}}(D_{\boldsymbol{d}}, D^*) \\ \mathbf{C}_{\mathsf{pr}}(D^*, D_{\boldsymbol{d}}) \mathbf{W}^T & \mathbf{C}_{\mathsf{pr}}(D^*, D^*) \end{bmatrix} \right)$$

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 \rightsquigarrow design-dependent posterior:

$$oldsymbol{m}^*|oldsymbol{d}(oldsymbol{w}) \sim \mathcal{N}(oldsymbol{m}_{\mathsf{post}}(oldsymbol{w}), \mathbf{C}_{\mathsf{post}}(oldsymbol{w}))$$

2. Description of "optimal" design

Goal: Choose measurement locations to minimize posterior "uncertainty"

▶ level of "uncertainty" measured by $\phi(w) := \phi(\mathbf{C}_{\mathsf{post}}(w))$

$$oldsymbol{w}^* = rgmin_{oldsymbol{w}\in\{0,1\}^s} \phi(oldsymbol{w})$$

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 $oldsymbol{w}^* = \operatorname*{arg\,min}_{oldsymbol{w}\in\{0,1\}^s} \phi(oldsymbol{w})$

► $\lambda_1(\boldsymbol{w}) \leq \lambda_2(\boldsymbol{w}) \leq \ldots \leq \lambda_n(\boldsymbol{w})$ eigenvalues of $\mathbf{C}_{\mathsf{post}}(\boldsymbol{w})$

Many choices for ϕ ...

A-optimal:
$$\phi^A(\boldsymbol{w}) = \operatorname{trace}\left[\mathbf{C}_{\mathsf{post}}(\boldsymbol{w})\right] = \sum_{i=1}^n \lambda_i(\boldsymbol{w})$$

D-optimal:
$$\phi^D(oldsymbol{w}) = \mathsf{det}\left[\mathbf{C}_{\mathsf{post}}(oldsymbol{w})
ight] = \prod_{i=1}^n \lambda_i(oldsymbol{w})$$

E-optimal: $\phi^{E}(\boldsymbol{w}) = \lambda_{n}(\mathbf{C}_{\mathsf{post}}(\boldsymbol{w}))$

$$oldsymbol{w}^* = rgmin_{oldsymbol{w}\in\{0,1\}^s} \phi^A(oldsymbol{w})$$

- Trivial solution: $w_i = 1$ for $i = 1, \ldots, s$
- Real-world applications, measurements often costly

$$oldsymbol{w}^* = rgmin_{oldsymbol{w}\in\{0,1\}^s} \phi^A(oldsymbol{w})$$

► Real-world applications, measurements often costly ⇒ introduce **cost constraints**:

$$egin{aligned} m{w}^*_{\mathsf{opt}} &= rgmin_{m{w}\in\{0,1\}^s} & \phi^A(m{w}) \ & \mathsf{s.t.} \ \sum_{i=1}^s w_i = k \end{aligned}$$

- ► Real-world applications, measurements often costly
- \implies introduce **cost constraints**:
 - 1. Direct combinatorial search \rightsquigarrow global optimal w^*_{opt}
 - ▶ Requires $\binom{s}{k}$ evaluations of $\phi^A(\boldsymbol{w}) = \text{trace}\left[\mathcal{C}_{\mathsf{post}}(\boldsymbol{w})\right]$

$$oldsymbol{w}_{\mathsf{G}}^* pprox oldsymbol{w}_{\mathsf{Opt}}^* = rgmin_{oldsymbol{w} \in \{0,1\}^s} \phi^A(oldsymbol{w})$$

s.t. $\sum_{i=1}^s w_i = k$

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 - 1. Direct combinatorial search \rightsquigarrow global optimal w^*_{opt}
 - Requires $\binom{s}{k}$ evaluations of $\phi^A(\boldsymbol{w}) = \operatorname{trace} [\mathcal{C}_{\mathsf{post}}(\boldsymbol{w})]$
 - 2. Greedy approach
 - \blacktriangleright Simple to implement, less ϕ^A evaluations but still many, suboptimal

$$oldsymbol{w}^* = rgmin_{oldsymbol{w}\in[0,1]^s} \phi^A(oldsymbol{w}) + \gamma\psi(oldsymbol{w})$$

- Real-world applications, measurements often costly
- \implies introduce **cost constraints**:
 - 1. Direct combinatorial search \rightsquigarrow global optimal w^*_{opt}
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 - 3. Relaxation + sparsification
 - # of ϕ^A evaluations does not grow with # of sensors, gradients of ϕ^A needed, indirect control of sparsity, suboptimal










Infinite-dimensional Bayesian inverse problems

Introduce non-trivial parameter-to-observable map $\mathcal{F}: \mathcal{H} \to \mathbb{R}^d$

$$d = \mathcal{F}(m) + \eta$$

► \mathcal{F} : PDE solve + spatiotemporal observation operator ► $m \sim \mu_0 = \mathcal{N}(0, \mathcal{C}_{pr}), \ \eta \sim \mathcal{N}(\mathbf{0}, \Gamma_{noise})$

Goal: Infer posterior measure for m given *indirect* noisy measurements d

Infinite-dimensionalBayesian inverse problemsIntroduce non-trivial parameter-to-observable map $\mathcal{F} : \mathcal{H} \to \mathbb{R}^d$

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Goal: Infer posterior measure for m given *indirect* noisy measurements dBayes' rule \rightsquigarrow posterior law on m:

$$\frac{d\mu_{\mathsf{post}}^{d}}{d\mu_{0}} \propto \pi_{\mathsf{like}}(d|m), \quad \pi_{\mathsf{like}}(d|m) \propto \exp\left[-\frac{1}{2}\|\mathcal{F}(m) - d\|_{\Gamma_{\mathsf{noise}}^{-1}}^{2}\right]$$

Infinite-dimensional *linear* Bayesian inverse problems Introduce non-trivial *parameter-to-observable map* $\mathcal{F} : \mathcal{H} \to \mathbb{R}^d$

$$d=\mathcal{F} \, \, m \, \, +\eta$$

► \mathcal{F} : PDE solve + spatiotemporal observation operator ► $m \sim \mu_0 = \mathcal{N}(0, \mathcal{C}_{\text{pr}}), \ \boldsymbol{\eta} \sim \mathcal{N}(\mathbf{0}, \boldsymbol{\Gamma}_{\text{noise}})$

Goal: Infer posterior measure for m given *indirect* noisy measurements dBayes' rule \rightsquigarrow posterior law on m:

$$\frac{d\mu_{\mathsf{post}}^{d}}{d\mu_{0}} \propto \pi_{\mathsf{like}}(\boldsymbol{d}|\boldsymbol{m}), \quad \pi_{\mathsf{like}}(\boldsymbol{d}|\boldsymbol{m}) \propto \exp\left[-\frac{1}{2}\|\mathcal{F} \ \boldsymbol{m} \ -\boldsymbol{d}\|_{\boldsymbol{\Gamma}_{\mathsf{noise}}^{-1}}^{2}\right]$$

▶ For linear \mathcal{F} , $m|d \sim \mathcal{N}(m_{\mathsf{post}}, \mathcal{C}_{\mathsf{post}})$ with

$$\mathcal{C}_{\mathsf{post}} = \left(\mathcal{F}^* \mathbf{\Gamma}_{\mathsf{noise}}^{-1} \mathcal{F} + \mathcal{C}_{\mathsf{pr}}^{-1}\right)^{-1}$$

OED for infinite-dimensional Bayesian inverse problems

As before:

- Grid of s possible sensor locations for measurement collection at r times $\implies d=rs$ observations
- ▶ Differentiate between designs through (block-)diagonal $\mathbf{W} \in \mathbb{R}^{d \times d}$

Other design definitions possible

Design enters through the likelihood:

$$\pi_{\mathsf{like}}(\boldsymbol{d}|m) \propto \exp\left[-rac{1}{2}\|\mathcal{F}m-\boldsymbol{d}\|_{\mathbf{\Gamma}_{\mathbf{W}}^{-1}}^{2}
ight]$$

- $\Gamma_{\mathbf{W}}^{-1}$ depends on noise model
- For uncorrelated noise, e.g., $\Gamma_{\text{noise}} = \sigma_n^2 \mathbf{I}$, $\Gamma_{\mathbf{W}}^{-1} := \frac{1}{\sigma_n^2} \mathbf{W}$

OED for infinite-dimensional Bayesian inverse problems

Design enters through the likelihood:

$$\pi_{\mathsf{like}}(\boldsymbol{d}|\boldsymbol{m}) \propto \exp\left[-rac{1}{2\sigma_n^2}\|\mathcal{F}\boldsymbol{m} - \boldsymbol{d}\|_{\mathbf{W}}^2
ight]$$

 \rightsquigarrow design-dependent posterior measure $m|d(w) \sim \mathcal{N}(m_{\text{post}}(w), \mathcal{C}_{\text{post}}(w))$

$$\mathcal{C}_{\mathsf{post}}(\boldsymbol{w}) = \left(\sigma_n^{-2}\mathcal{F}^*\mathbf{W}\mathcal{F} + \mathcal{C}_{\mathsf{pr}}^{-1}\right)^{-1}$$

Infinite-dimensional A-optimality criterion and challenges

► Infinite-dimensional A-optimality criterion defined by:

$$\phi^{A}(\boldsymbol{w}) = \operatorname{trace}\left[\mathcal{C}_{\mathsf{post}}(\boldsymbol{w})\right] = \operatorname{trace}\left[\left(\sigma_{n}^{-2}\mathcal{F}^{*}\mathbf{W}\mathcal{F} + \mathcal{C}_{\mathsf{pr}}^{-1}\right)^{-1}\right]$$

Finding A-optimal designs is challenging:

- \blacktriangleright Requires many evaluations of ϕ^A
- Computing trace of ∞ -dimensional, PDE-dependent operator
- Finding global or greedy optimal is too expensive

1. Approximate the trace

- Use randomized trace estimation
- Alternatively: reformulate ϕ^A to reduce dimensionality

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2. Eliminate PDEs

- \blacktriangleright Exploit low-rank structure of ${\cal F}$
- \blacktriangleright Approximate ${\mathcal F}$ with truncated SVD using matrix-free algorithms

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- Use randomized trace estimation
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2. Eliminate PDEs

- Exploit low-rank structure of \mathcal{F}
- Approximate \mathcal{F} with truncated SVD using matrix-free algorithms
- 3. Enforce sparse designs with sparsity-inducing penalty $\boldsymbol{\psi}$

$$oldsymbol{w}^* = rgmin_{oldsymbol{w}\in[0,1]^s} \phi^A(oldsymbol{w}) + \gamma\psi \quad (oldsymbol{w})$$

▶ $\psi(\boldsymbol{w}) \approx \|\boldsymbol{w}\|_0$, the number of non-zero weights

1. Approximate the trace

- Use randomized trace estimation
- Alternatively: reformulate ϕ^A to reduce dimensionality

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- \blacktriangleright Exploit low-rank structure of ${\cal F}$
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- 3. Enforce sparse designs with sparsity-inducing penalty $\psi_{arepsilon(i)}$

$$oldsymbol{w}^*_{arepsilon(i)} = rgmin_{oldsymbol{w}\in[0,1]^s} \phi^A(oldsymbol{w}) + \gamma \psi_{arepsilon(i)}(oldsymbol{w})$$

▶ $\psi(\boldsymbol{w}) \approx \|\boldsymbol{w}\|_0$, the number of non-zero weights

▶ ℓ_0 -sparsification, $\psi_{\varepsilon(i)} \to \| \cdot \|_0$ as $i \to \infty$

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Motivating example

Many models for real-world phenomena have uncertain inputs

Ex: Contaminant source identification in groundwater flow:



Designs need to work well for all realizations of uncertainty



OED under uncertainty

Common OED assumptions:

- exact knowledge of model equations
- no other sources of uncertainty

Aim of this work:

- Formulation of OED under irreducible uncertainty
- Mathematical structure and computational challenges

OED for Bayesian linear inverse problems under uncertainty Incorporate uncertainty into $\mathcal{F}, \mathcal{F} : (\Omega, \mathcal{G}, P) \to \mathcal{L}(\mathcal{H}, \mathbb{R}^d)$

 $d(\xi) = \mathcal{F}(\xi)m + \eta$

Likelihood depends on uncertainty:

$$\pi_{\mathsf{like}}(\boldsymbol{d}|m) \propto \exp\left[-\frac{1}{2\sigma_n^2} \|\mathcal{F}(\xi)m - \boldsymbol{d}\|_{\mathbf{W}}^2
ight]$$

 \implies posterior depends on uncertainty:

$$\mu_{\mathsf{post}}^{d} = \mathcal{N}(m_{\mathsf{post}}(\xi, \boldsymbol{w}), \mathcal{C}_{\mathsf{post}}(\xi, \boldsymbol{w}))$$

$$\mathcal{C}_{\mathsf{post}}(\xi, \boldsymbol{w}) = \left(\frac{1}{\sigma_n^2} \mathcal{F}^*(\xi) \mathbf{W} \mathcal{F}(\xi) + \mathcal{C}_{\mathsf{pr}}^{-1}\right)^{-1}$$

A-optimal design under uncertainty

A-optimal design under uncertainty:

$$\boldsymbol{w}^{*} = \underset{\boldsymbol{w} \in [0,1]^{d}}{\operatorname{arg\,min}} \int_{\Omega} \operatorname{trace} \left[\mathcal{C}_{\mathsf{post}}(\xi, \boldsymbol{w}) \right] P(d\xi) + \gamma \psi\left(\boldsymbol{w} \right)$$

- Minimizes expected value of average posterior variance
- \blacktriangleright Uncertainty-aware designs do well on average, but are not optimal given fixed ξ

Computational challenges

- 1. Discretization of uncertainty
- 2. Efficient computation of trace
- 3. Tractable computation of optimal designs

1. Discretization of the uncertainty

- Approximate the expected value of the average pointwise posterior variance
- Assuming we can sample $\xi_i \in \Omega$, we use SAA to approximate the integral:

$$\int_{\Omega} \operatorname{trace} \left[\mathcal{C}_{\mathsf{post}}(\xi, \boldsymbol{w}) \right] P(d\xi) \approx \frac{1}{N} \sum_{i=1}^{N} \operatorname{trace} \left[\mathcal{C}_{\mathsf{post}}(\xi_i, \boldsymbol{w}) \right]$$

2. Computation of trace, "measurement space approach" Discretized trace (using, e.g., finite elements, $\mathcal{F} \approx \mathbf{F} \in \mathbb{R}^{d \times n}$):

$$\phi^A(\xi, \boldsymbol{w}) \approx \phi_n^A(\xi, \boldsymbol{w}) = \operatorname{trace}\left[\left(\frac{1}{\sigma_n^2} \mathbf{F}(\xi)^* \mathbf{W} \mathbf{F}(\xi) + \mathbf{C}_{\mathsf{pr}}^{-1}\right)^{-1}\right]$$

Too expensive to compute trace exactly even after discretization
 We can rewrite φ^A_n(ξ, w) as:

$$\phi_n^A(\xi, \boldsymbol{w}) = \operatorname{trace} \left[\mathbf{C}_{\mathsf{pr}} \right] - \operatorname{trace} \left[\frac{1}{\sigma_n^2} \mathbf{S}^{-1}(\xi, \boldsymbol{w}) \mathbf{W} \mathbf{F}(\xi) \mathbf{C}_{\mathsf{pr}}^2 \mathbf{F}^*(\xi) \right]$$
$$= \operatorname{trace} \left[\mathbf{C}_{\mathsf{pr}} \right] - \operatorname{trace} \left[\mathbf{K}(\xi, \boldsymbol{w}) \right]$$

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Too expensive to compute trace exactly even after discretization
 We can rewrite φ^A_n(ξ, w) as:

$$\phi_n^A(\xi, \boldsymbol{w}) = \operatorname{trace} \left[\mathbf{C}_{\mathsf{pr}} \right] - \operatorname{trace} \left[\frac{1}{\sigma_n^2} \mathbf{S}^{-1}(\xi, \boldsymbol{w}) \mathbf{W} \mathbf{F}(\xi) \mathbf{C}_{\mathsf{pr}}^2 \mathbf{F}^*(\xi) \right]$$
$$= \operatorname{trace} \left[\mathbf{C}_{\mathsf{pr}} \right] - \operatorname{trace} \left[\mathbf{K}(\xi, \boldsymbol{w}) \right]$$

Optimal design satisfies:

$$\boldsymbol{w}^* = \operatorname*{arg\,min}_{\boldsymbol{w} \in [0,1]^d} \left[-\frac{1}{N} \sum_{i=1}^N \operatorname{trace} \left[\mathbf{K}(\xi_i, \boldsymbol{w}) \right] + \gamma \psi \left(\boldsymbol{w} \right) \right]$$

Trace of an operator in measurement space (finite)

3. Elimination of PDEs from the minimization

OEDUU objective is expensive to optimize:

$$\frac{1}{N}\sum_{i=1}^{N}\operatorname{trace}\left[\mathbf{K}(\xi_{i},\boldsymbol{w})\right]$$

- $\mathbf{K}(\xi_i, \boldsymbol{w})$ depends on \mathbf{F}_i and \mathbf{F}_i^*
- Computing trace even for one ξ_i requires many PDE solves
- Need to compute trace for each sample ξ_i

3. Elimination of PDEs from the minimization

Find a low-rank approximation to $\mathbf{F}(\xi_i)\mathbf{C}_{pr}^{rac{1}{2}} = \widetilde{\mathbf{F}}_i$

- Preconditioning promotes faster decay of eigenvalues
- Matrix-free techniques based on randomized linear algebra

Storing separate basis vectors for each $\widetilde{\mathbf{F}}_i$ is infeasible

► Solution: find a space that captures the "effective" composite range space for all F(ξ_i)

Find $\mathbf{Q} \in \mathbb{R}^{d \times k}$ and $\widehat{\mathbf{Q}} \in \mathbb{R}^{m \times k}$ (k small) such that $\forall i \in \{1, \dots, N\}$: $\widetilde{\mathbf{F}}_i \approx \mathbf{Q} \mathbf{Q}^* \widetilde{\mathbf{F}}_i \widehat{\mathbf{Q}} \widehat{\mathbf{Q}}^*$

Many ways to make this more efficient...

Numerical example - subsurface flow OED



$$\begin{split} u_t - \kappa \Delta u + v(\xi) \cdot \nabla u &= 0 & \text{ in } \mathcal{D} \times (T_i, T) \\ u(\cdot, T_i) &= m & \text{ in } \mathcal{D} \\ -\kappa \nabla u \cdot n + v(\xi) \cdot nu &= 0 & \text{ on } \Gamma_L \times (T_i, T) \\ \kappa \nabla u \cdot n &= 0 & \text{ on } \Gamma_O \times (T_i, T) \end{split}$$

▶ Grid of 234 sensor locations, measurements taken at T_j ∈ {τ₁,...,τ_r}
 ▶ Samples {v(ξ_i)}^N_{i=1} of the velocity field and T_i ~ U[-1,1] of initial time

~> Find subset of locations minimizing A-optimal criterion under uncertainty

Spatial/temporal discretization: built on FEniCS and hIPPYlib (open source Python/C++ framework)

Subsurface flow OEDUU

$$oldsymbol{w}^* = rgmin_{oldsymbol{w}\in[0,1]^d} \left[-rac{1}{N} \sum_{i=1}^N ext{trace} \left[\mathbf{K}(\xi_i, oldsymbol{w})
ight] + \gamma \psi(oldsymbol{w})
ight]$$

- $\blacktriangleright~N=100$ samples for discretization of uncertainty
- ℓ_0 -sparsification used to find sparse designs
- Each minimization solved with gradient-based method (projected BFGS)

Deterministic vs. designs under uncertainty



Deterministic vs. designs under uncertainty



Motivating examples

Introduction and background OED for Gaussian regression OED for Bayesian inverse problem:

OED under model uncertainty Mathematical formulation of OED Computational challenges Numerical results – subsurface flow

OED for tsunami source reconstruction Mathematical formulation Numerical results

Summary

Earthquake-generated tsunamis



- Tsunamis generated by earthquakes beneath ocean floor at subduction zones
- Water pressure/height readings are used to detect and track tsunamis (DART system)
- Tsunami detection and warning relies on informative data

Earthquake-generated tsunamis

Governing equation for \mathcal{G} :

$$\begin{bmatrix} h \\ u \\ v \end{bmatrix}_t + \begin{bmatrix} u \\ \frac{u^2}{h} + \frac{1}{2}gh^2 \\ \frac{uv}{h} \end{bmatrix}_x + \begin{bmatrix} v \\ \frac{uv}{h} \\ \frac{v^2}{h} + \frac{1}{2}gh^2 \end{bmatrix}_y = - \begin{bmatrix} 0 \\ ghB_x \\ ghB_y \end{bmatrix}$$

Goal: Find optimal configuration of sensors for inference of B



Gaussian approximation to posterior distribution SWE nonlinear \implies

- 1. solutions can exhibit shocks
- 2. non-Gaussian posterior

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Solutions well-approximated by linearization in deep water:

$$\mathcal{G}(B) \approx \mathcal{G}(B_0) + \mathcal{F}[B - B_0], \quad \mathcal{F} := \mathcal{G}'(B_0)$$

 $\label{eq:linearization} \begin{array}{l} \underset{}{\overset{}{\overset{}}{\underset{}}} \text{ Gaussian approximation to posterior:} \\ B| \boldsymbol{d} \sim \mathcal{N}(B_{\mathsf{post}}(\boldsymbol{w}), \mathcal{C}_{\mathsf{post}}(\boldsymbol{w})) \text{ with} \end{array}$

$$\mathcal{C}_{\mathsf{post}}(oldsymbol{w}) = \left(rac{1}{\sigma_n^2}\mathcal{F}^*\mathbf{W}\mathcal{F} + \mathcal{C}_{\mathsf{pr}}^{-1}
ight)^{-1}$$

Prior on B

Bathymetry change $\hat{B} := B - B_0$ is due to a slip at a fault

► Okada model ~→ linear relationship between slips S at m slip patches and seafloor deformation

$\hat{B} = O\mathbf{S}$

▶ Prior on slips $\mathbf{S} \sim \mathcal{N}(\mathbf{0}, \theta^2 \mathbf{I})$ induces prior on $B \sim \mathcal{N}(B_0, \mathcal{C}_{pr}(O))$

Reasonable sample seafloor deformations



Inversion for slips S

Exploiting linear relationship between ${\bf S}$ and $\hat{B} \rightsquigarrow$

Reformulation of inverse problem:

$$\boldsymbol{d} = \mathcal{G}(B_0) + \mathcal{F}\hat{B} + \boldsymbol{\eta} = \mathcal{G}(B_0) + \mathcal{F}O\mathbf{S} + \boldsymbol{\eta}$$

 \rightsquigarrow *finite-dimensional* posterior distribution for slips $\mathbf{S} \in \mathbb{R}^m$:

$$\mathbf{S}|m{d} \sim \mathcal{N}(\mathbf{S}_{\mathsf{post}}(m{w}), \mathbf{C}_{\mathsf{post}}(m{w}))$$
 with

$$\mathbf{C}_{\mathsf{post}}(\boldsymbol{w}) = \left(\frac{1}{\sigma_n^2} O^* \mathcal{F}^* \mathbf{W} \mathcal{F} O + \theta^{-2} \mathbf{I}\right)^{-1}$$

2D example - problem setup



- ▶ s = 189 possible locations
- r = 8 observation times
- ▶ m = 20 slip patches
- GeoClaw used for numerical results

2D example - OED problem

$$\boldsymbol{w}_{\mathsf{G}}^* \approx \operatorname*{arg\,min}_{\substack{\boldsymbol{w} \in \{0,1\}^s \\ \text{s.t.} \sum_{i=1}^s \boldsymbol{w}_i = k}} \operatorname{trace} \left[\left(\frac{1}{\sigma_n^2} (\mathcal{FO})^* \mathbf{W}(\mathcal{FO}) + \theta^{-2} \mathbf{I} \right)^{-1} \right]$$

- $\blacktriangleright \ \mathcal{FO} \in \mathbb{R}^{d \times m}$ precomputed
- No adjoint solves needed
- Greedy optimal weight vectors computed, PDE free





2D example - design comparisons


2D example - design comparisons



Motivating examples

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Summary

Summary and future work

Summary:

- 1. OED under uncertainty:
- Introduced mathematical framework for incorporation of uncertainty
- Presented "measurement space approach" formulation of OED objective
- Eliminated PDEs from minimization using *joint* basis
- Demonstrated effectiveness of OEDUU using numerical example

Summary and future work

Summary:

- $1. \ \mathsf{OED} \ \mathsf{under} \ \mathsf{uncertainty:}$
- Introduced mathematical framework for incorporation of uncertainty
- Presented "measurement space approach" formulation of OED objective
- Eliminated PDEs from minimization using *joint* basis
- Demonstrated effectiveness of OEDUU using numerical example
- 2. OED for tsunami source reconstruction:
- Formulated OED problem for deep-ocean tsunami source reconstruction using SWE
- Used Gaussian approximation to posterior through linearization
- Reformulated problem to invert for slips allowing elimination of PDEs

Summary and future work

Possible extensions:

- Alternate ways of dealing with uncertainty, e.g., stochastic approximation or Taylor expansion
- Laplace approximation to posterior
- Inclusion of uncertain parameters into tsunami model
- Incorporate OED framework into GeoClaw

Thank you!

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