

Extreme Eigenvalues of Random Matrices

Satya N. Majumdar

Laboratoire de Physique Théorique et Modèles Statistiques, CNRS,
Université Paris-Sud, France

June 15, 2009

Plan

Plan:

- The problem & its origin

Plan

Plan:

- The problem & its origin
- Top eigenvalue λ_{\max} of a Gaussian random matrix and its large deviations

Plan:

- The problem & its origin
- Top eigenvalue λ_{\max} of a Gaussian random matrix and its large deviations
- Prob. distr. of $\lambda_{\max} \iff$ Coulomb gas with a wall

Plan:

- The problem & its origin
- Top eigenvalue λ_{\max} of a Gaussian random matrix and its large deviations
- Prob. distr. of $\lambda_{\max} \iff$ Coulomb gas with a wall
- Extension to Wishart matrices (Laguerre ensemble)

Plan:

- The problem & its origin
- Top eigenvalue λ_{\max} of a Gaussian random matrix and its large deviations
- Prob. distr. of $\lambda_{\max} \iff$ Coulomb gas with a wall
- Extension to Wishart matrices (Laguerre ensemble)
- Summary and Conclusions

A Trivial Problem

DIAGONAL MATRIX

$$\begin{bmatrix} X_{11} & & & & & & \\ & X_{22} & & & & & \\ & & X_{33} & & & & \\ & & & \ddots & & & \\ & & & & \ddots & & \\ & & & & & \ddots & \\ & & & & & & X_{NN} \end{bmatrix}$$

$$\begin{aligned} \Pr[X_{ii} = x] \\ = (2\pi)^{-1/2} \exp[-x^2/2] \end{aligned}$$



GAUSSIAN

N Eigenvalues: $\lambda_i = X_{ii} \rightarrow$ Independent

A Trivial Problem

DIAGONAL MATRIX

$$\begin{bmatrix} X_{11} & & & & \\ & X_{22} & & & \\ & & X_{33} & & \\ & & & \ddots & \\ 0 & & & & X_{NN} \end{bmatrix} \quad \begin{array}{l} \Pr[X_{ii} = x] \\ = (2\pi)^{-1/2} \exp[-x^2/2] \\ \downarrow \\ \text{GAUSSIAN} \end{array}$$

N Eigenvalues: $\lambda_i = X_{ii} \rightarrow$ Independent

- $P_N = \text{Prob}[\lambda_1 \leq 0, \lambda_2 \leq 0, \dots, \lambda_N \leq 0] = 2^{-N} = \exp[-(\ln 2) N]$

A Nontrivial Problem

REAL SYMMETRIC MATRIX (N×N)

$$\mathbf{X} = \begin{pmatrix} x_{11} & x_{12} & \cdots & \cdots & x_{1N} \\ x_{21} & x_{22} & \cdots & \cdots & x_{2N} \\ \vdots & \vdots & & & \vdots \\ \vdots & \vdots & & & \vdots \\ \vdots & \vdots & & & \vdots \\ x_{N1} & \cdots & \cdots & \cdots & x_{NN} \end{pmatrix}$$

GAUSSIAN
↓
Pr [X]
∝
 $\exp[-\frac{1}{2}\text{Tr}(x^2)]$

N eigenvalues : $\lambda_1, \lambda_2, \dots, \lambda_N$
↳ strongly correlated

A Nontrivial Problem

REAL SYMMETRIC MATRIX (N x N)

$$\mathbf{X} = \begin{pmatrix} x_{11} & x_{12} & \cdots & \cdots & x_{1N} \\ x_{21} & x_{22} & \cdots & \cdots & x_{2N} \\ \vdots & \vdots & & & \vdots \\ \vdots & \vdots & & & \vdots \\ \vdots & \vdots & & & \vdots \\ x_{N1} & \cdots & \cdots & \cdots & x_{NN} \end{pmatrix}$$

GAUSSIAN
↓
Pr [X]
∝
 $\exp[-\frac{1}{2}\text{Tr}(x^2)]$

N eigenvalues : $\lambda_1, \lambda_2, \dots, \lambda_N$
↳ strongly correlated

- $P_N = \text{Prob}[\lambda_1 \leq 0, \lambda_2 \leq 0, \dots, \lambda_N \leq 0] = \text{Prob}[\lambda_{\max} \leq 0] = ?$

A Nontrivial Problem

REAL SYMMETRIC MATRIX (N×N)

$$\mathbf{X} = \begin{pmatrix} x_{11} & x_{12} & \cdots & \cdots & x_{1N} \\ x_{21} & x_{22} & \cdots & \cdots & x_{2N} \\ \vdots & \vdots & & & \vdots \\ \vdots & \vdots & & & \vdots \\ \vdots & \vdots & & & \vdots \\ x_{N1} & \cdots & \cdots & \cdots & x_{NN} \end{pmatrix}$$

GAUSSIAN
↓
Pr [X]
∝
 $\exp[-\frac{1}{2}\text{Tr}(\mathbf{X}^2)]$

N eigenvalues : $\lambda_1, \lambda_2, \dots, \lambda_N$
↳ strongly correlated

- $P_N = \text{Prob}[\lambda_1 \leq 0, \lambda_2 \leq 0, \dots, \lambda_N \leq 0] = \text{Prob}[\lambda_{\max} \leq 0] = ?$

[R.M. May, Nature, 238, 413 (1972)—Ecosystems]

[Cavagna et. al. 2000, Fyodorov 2004, — Glassy systems]

[Suskind 2003, Douglas et. al. 2004, Aazami & Easter 2006— String theory].....

Results for P_N :

- $P_N = \text{Prob}[\lambda_1 \leq 0, \lambda_2 \leq 0, \dots, \lambda_N \leq 0] = ?$

Results for P_N :

- $P_N = \text{Prob}[\lambda_1 \leq 0, \lambda_2 \leq 0, \dots, \lambda_N \leq 0] = ?$
- $N = 1$: $P_1 = 1/2 = 0.5$ (trivially)

Results for P_N :

- $P_N = \text{Prob}[\lambda_1 \leq 0, \lambda_2 \leq 0, \dots, \lambda_N \leq 0] = ?$
- $N = 1$: $P_1 = 1/2 = 0.5$ (trivially)
- $N = 2$: $P_2 = \frac{2-\sqrt{2}}{4} = 0.146447..$

Results for P_N :

- $P_N = \text{Prob}[\lambda_1 \leq 0, \lambda_2 \leq 0, \dots, \lambda_N \leq 0] = ?$
- $N = 1$: $P_1 = 1/2 = 0.5$ (trivially)
- $N = 2$: $P_2 = \frac{2-\sqrt{2}}{4} = 0.146447..$
- $N = 3$: $P_3 = \frac{\pi-2\sqrt{2}}{4\pi} = 0.0249209..$
(Beltrani 2007, Dedieu & Malajovich, 2007)

Results for P_N :

- $P_N = \text{Prob}[\lambda_1 \leq 0, \lambda_2 \leq 0, \dots, \lambda_N \leq 0] = ?$
- $N = 1$: $P_1 = 1/2 = 0.5$ (trivially)
- $N = 2$: $P_2 = \frac{2-\sqrt{2}}{4} = 0.146447..$
- $N = 3$: $P_3 = \frac{\pi-2\sqrt{2}}{4\pi} = 0.0249209..$
(Beltrani 2007, Dedieu & Malajovich, 2007)

Question: How does P_N decay for large N , i.e., $P_N \rightarrow ?$ as $N \rightarrow \infty$

Results for P_N :

- $P_N = \text{Prob}[\lambda_1 \leq 0, \lambda_2 \leq 0, \dots, \lambda_N \leq 0] = ?$

- $N = 1$: $P_1 = 1/2 = 0.5$ (trivially)

- $N = 2$: $P_2 = \frac{2-\sqrt{2}}{4} = 0.146447..$

- $N = 3$: $P_3 = \frac{\pi-2\sqrt{2}}{4\pi} = 0.0249209..$

(Beltrani 2007, Dedieu & Malajovich, 2007)

Question: How does P_N decay for large N , i.e., $P_N \rightarrow ?$ as $N \rightarrow \infty$

- Based on numerics, Aazami & Easter (2006) predicted for large N :

$$P_N \sim \exp[-\theta N^2] \text{ with } \theta_{\text{num}} \approx 0.27$$

→ very small probability → **RARE EVENT**

Results for P_N :

- $P_N = \text{Prob}[\lambda_1 \leq 0, \lambda_2 \leq 0, \dots, \lambda_N \leq 0] = ?$

- $N = 1$: $P_1 = 1/2 = 0.5$ (trivially)

- $N = 2$: $P_2 = \frac{2-\sqrt{2}}{4} = 0.146447..$

- $N = 3$: $P_3 = \frac{\pi-2\sqrt{2}}{4\pi} = 0.0249209..$

(Beltrani 2007, Dedieu & Malajovich, 2007)

Question: How does P_N decay for large N , i.e., $P_N \rightarrow ?$ as $N \rightarrow \infty$

- Based on numerics, Aazami & Easter (2006) predicted for large N :

$$P_N \sim \exp[-\theta N^2] \text{ with } \theta_{\text{num}} \approx 0.27$$

→ very small probability → **RARE EVENT**

- **Exact** result: $\theta = \frac{1}{4} \ln(3) = 0.274653..$ (Dean and S.M., 2006)

Results for P_N :

- $P_N = \text{Prob}[\lambda_1 \leq 0, \lambda_2 \leq 0, \dots, \lambda_N \leq 0] = ?$

- $N = 1$: $P_1 = 1/2 = 0.5$ (trivially)

- $N = 2$: $P_2 = \frac{2-\sqrt{2}}{4} = 0.146447..$

- $N = 3$: $P_3 = \frac{\pi-2\sqrt{2}}{4\pi} = 0.0249209..$

(Beltrani 2007, Dedieu & Malajovich, 2007)

Question: How does P_N decay for large N , i.e., $P_N \rightarrow ?$ as $N \rightarrow \infty$

- Based on numerics, Aazami & Easter (2006) predicted for large N :

$$P_N \sim \exp[-\theta N^2] \text{ with } \theta_{\text{num}} \approx 0.27$$

→ very small probability → **RARE EVENT**

- **Exact** result: $\theta = \frac{1}{4} \ln(3) = 0.274653..$ (Dean and S.M., 2006)

More generally, for $\beta = 1$ (**GOE**), $\beta = 2$ (**GUE**) and $\beta = 4$ (**GSE**)

$$P_N \sim \exp[-\beta\theta N^2] \text{ for large } N$$

Complex Landscapes

A particle moving in a N -dimensional landscape: $V(y_1, y_2, \dots, y_N)$

$$\frac{dy_i}{dt} = -\nabla_{y_i} V$$

Spin glasses, Structural glasses, Supercooled liquids, String landscapes

Complex Landscapes

A particle moving in a N -dimensional landscape: $V(y_1, y_2, \dots, y_N)$

$$\frac{dy_i}{dt} = -\nabla_{y_i} V$$

Spin glasses, Structural glasses, Supercooled liquids, String landscapes

- Stationary point (fixed pt.) y^* : $\nabla V = 0$ at $y = y^*$

Complex Landscapes

A particle moving in a N -dimensional landscape: $V(y_1, y_2, \dots, y_N)$

$$\frac{dy_i}{dt} = -\nabla_{y_i} V$$

Spin glasses, Structural glasses, Supercooled liquids, String landscapes

- **Stationary** point (fixed pt.) y^* : $\nabla V = 0$ at $y = y^*$

- Near a stationary point: $V(\{y_i\}) \approx \sum_{i,j} H_{i,j}(y_i - y_i^*)(y_j - y_j^*)$

Hessian matrix: $H_{i,j} \equiv \left[\frac{\partial^2 V}{\partial y_i \partial y_j} \right]$

Complex Landscapes

A particle moving in a N -dimensional landscape: $V(y_1, y_2, \dots, y_N)$

$$\frac{dy_i}{dt} = -\nabla_{y_i} V$$

Spin glasses, Structural glasses, Supercooled liquids, String landscapes

- **Stationary** point (fixed pt.) y^* : $\nabla V = 0$ at $y = y^*$

- Near a stationary point: $V(\{y_i\}) \approx \sum_{i,j} H_{i,j}(y_i - y_i^*)(y_j - y_j^*)$

$$\text{Hessian matrix: } H_{i,j} \equiv \left[\frac{\partial^2 V}{\partial y_i \partial y_j} \right]$$

- **Eigenvalues** of the **Hessian** (stability) matrix determines the **nature** of the stationary point:

\implies **Local Minimum, Local Maximum & Saddles**

Eigenvalues of the Hessian Matrix

Examples:

- $N = 1$ -dimensional surface: Hessian matrix $H = \frac{\partial^2 V}{\partial y^2}$

If $\frac{\partial^2 V}{\partial y^2} < 0 \rightarrow$ Local Maximum; if $\frac{\partial^2 V}{\partial y^2} > 0 \rightarrow$ Local Minimum

Eigenvalues of the Hessian Matrix

Examples:

- $N = 1$ -dimensional surface: Hessian matrix $H = \frac{\partial^2 V}{\partial y^2}$

If $\frac{\partial^2 V}{\partial y^2} < 0 \rightarrow$ Local Maximum; if $\frac{\partial^2 V}{\partial y^2} > 0 \rightarrow$ Local Minimum

- $N = 2$ -dimensional surface: Hessian matrix $H \equiv \begin{pmatrix} \frac{\partial^2 V}{\partial y_1^2} & \frac{\partial^2 V}{\partial y_1 \partial y_2} \\ \frac{\partial^2 V}{\partial y_2 \partial y_1} & \frac{\partial^2 V}{\partial y_2^2} \end{pmatrix}$

Eigenvalues of the Hessian Matrix

Examples:

- $N = 1$ -dimensional surface: Hessian matrix $H = \frac{\partial^2 V}{\partial y^2}$

If $\frac{\partial^2 V}{\partial y^2} < 0 \rightarrow$ Local Maximum; if $\frac{\partial^2 V}{\partial y^2} > 0 \rightarrow$ Local Minimum

- $N = 2$ -dimensional surface: Hessian matrix $H \equiv \begin{pmatrix} \frac{\partial^2 V}{\partial y_1^2} & \frac{\partial^2 V}{\partial y_1 \partial y_2} \\ \frac{\partial^2 V}{\partial y_2 \partial y_1} & \frac{\partial^2 V}{\partial y_2^2} \end{pmatrix}$

Two real eigenvalues: (λ_1, λ_2)

If $\lambda_1 < 0$ and $\lambda_2 < 0 \rightarrow$ Local Maximum

If $\lambda_1 > 0$ and $\lambda_2 > 0 \rightarrow$ Local Minimum

$\left. \begin{array}{l} \lambda_1 < 0, \quad \lambda_2 > 0 \\ \lambda_1 > 0, \quad \lambda_2 < 0 \end{array} \right\} \rightarrow$ Saddle

Fraction of Local Maximum/Minimum in Random Hessian Model

- Random $(N \times N)$ Hessian Model: $H \equiv \left[\frac{\partial^2 V}{\partial y_i \partial y_j} \right]$
 $H \rightarrow (N \times N)$ real symmetric random matrix

Fraction of Local Maximum/Minimum in Random Hessian Model

- Random $(N \times N)$ Hessian Model: $H \equiv \left[\frac{\partial^2 V}{\partial y_i \partial y_j} \right]$
 $H \rightarrow (N \times N)$ real symmetric random matrix
- If the entries are Gaussian: $H \rightarrow (\text{GOE})$
[Cavagna et. al. (2000), ..., Aazami & Easter (2006)]

Fraction of Local Maximum/Minimum in Random Hessian Model

- Random $(N \times N)$ Hessian Model: $H \equiv \left[\frac{\partial^2 V}{\partial y_i \partial y_j} \right]$
 $H \rightarrow (N \times N)$ real symmetric random matrix
- If the entries are Gaussian: $H \rightarrow (\text{GOE})$
[Cavagna et. al. (2000), ..., Aazami & Easter (2006)]
- In the random Hessian model

$P_N =$ Fraction of Local Maxima/minima

$$= \text{Prob}[\lambda_1 < 0, \lambda_2 < 0, \dots, \lambda_N < 0]$$

$$= \text{Prob}[\lambda_{\max} < 0]$$

$$\xrightarrow{N \rightarrow \infty} \exp[-\theta N^2] \rightarrow \text{very small}$$

Fraction of Local Maximum/Minimum in Random Hessian Model

- Random ($N \times N$) Hessian Model: $H \equiv \left[\frac{\partial^2 V}{\partial y_i \partial y_j} \right]$
 $H \rightarrow (N \times N)$ real symmetric random matrix
- If the entries are Gaussian: $H \rightarrow (\text{GOE})$
[Cavagna et. al. (2000), ..., Aazami & Easter (2006)]
- In the random Hessian model
 $P_N =$ Fraction of Local Maxima/minima
 $= \text{Prob}[\lambda_1 < 0, \lambda_2 < 0, \dots, \lambda_N < 0]$
 $= \text{Prob}[\lambda_{\max} < 0]$
 $\xrightarrow{N \rightarrow \infty} \exp[-\theta N^2] \rightarrow \text{very small}$
- \rightarrow Most of the stationary points \rightarrow Saddles

Gaussian Random Matrices

- $(N \times N)$ Gaussian random matrix: $X \equiv [x_{i,j}]$

Gaussian Random Matrices

- $(N \times N)$ Gaussian random matrix: $X \equiv [x_{i,j}]$
- Ensembles: Orthogonal (GOE), Unitary (GUE) or Symplectic (GSE)
- $\text{Prob}[x_{i,j}] \propto \exp \left[-\frac{\beta}{2} \text{Tr}(X, X) \right]$
where the Dyson index $\beta = 1$ (GOE), $\beta = 2$ (GUE) or $\beta = 4$ (GSE)

Gaussian Random Matrices

- $(N \times N)$ Gaussian random matrix: $X \equiv [x_{i,j}]$
- Ensembles: Orthogonal (GOE), Unitary (GUE) or Symplectic (GSE)
- $\text{Prob}[x_{i,j}] \propto \exp \left[-\frac{\beta}{2} \text{Tr}(X, X) \right]$
where the Dyson index $\beta = 1$ (GOE), $\beta = 2$ (GUE) or $\beta = 4$ (GSE)
- N real eigenvalues $\{\lambda_1, \lambda_2, \dots, \lambda_N\} \rightarrow$ correlated random variables

Gaussian Random Matrices

- $(N \times N)$ Gaussian random matrix: $X \equiv [x_{i,j}]$
- Ensembles: Orthogonal (GOE), Unitary (GUE) or Symplectic (GSE)
- $\text{Prob}[x_{i,j}] \propto \exp \left[-\frac{\beta}{2} \text{Tr}(X, X) \right]$
where the Dyson index $\beta = 1$ (GOE), $\beta = 2$ (GUE) or $\beta = 4$ (GSE)
- N real eigenvalues $\{\lambda_1, \lambda_2, \dots, \lambda_N\} \rightarrow$ correlated random variables
- Joint distribution of eigenvalues (Wigner, 1951)

$$P(\lambda_1, \lambda_2, \dots, \lambda_N) = \frac{1}{Z_N} \exp \left[-\frac{\beta}{2} \sum_{i=1}^N \lambda_i^2 \right] \prod_{j < k} |\lambda_j - \lambda_k|^\beta$$

Gaussian Random Matrices

- $(N \times N)$ Gaussian random matrix: $X \equiv [x_{i,j}]$
- Ensembles: Orthogonal (GOE), Unitary (GUE) or Symplectic (GSE)
- $\text{Prob}[x_{i,j}] \propto \exp \left[-\frac{\beta}{2} \text{Tr}(X, X) \right]$
where the Dyson index $\beta = 1$ (GOE), $\beta = 2$ (GUE) or $\beta = 4$ (GSE)
- N real eigenvalues $\{\lambda_1, \lambda_2, \dots, \lambda_N\} \rightarrow$ correlated random variables
- Joint distribution of eigenvalues (Wigner, 1951)

$$P(\lambda_1, \lambda_2, \dots, \lambda_N) = \frac{1}{Z_N} \exp \left[-\frac{\beta}{2} \sum_{i=1}^N \lambda_i^2 \right] \prod_{j < k} |\lambda_j - \lambda_k|^\beta$$

- where $Z_N =$ Partition Function

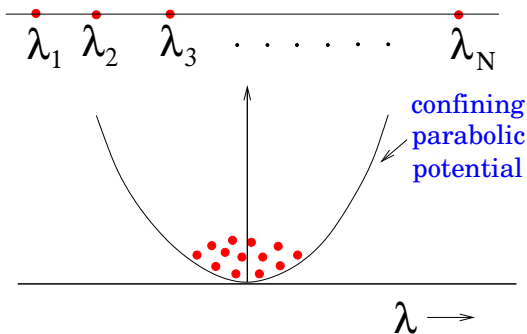
$$= \int_{-\infty}^{\infty} \dots \int_{-\infty}^{\infty} \left\{ \prod_i d\lambda_i \right\} \exp \left[-\frac{\beta}{2} \sum_{i=1}^N \lambda_i^2 \right] \prod_{j < k} |\lambda_j - \lambda_k|^\beta$$

Coulomb Gas interpretation

- $Z_N = \int_{-\infty}^{\infty} \dots \int_{-\infty}^{\infty} \left\{ \prod_i d\lambda_i \right\} \exp \left[-\frac{\beta}{2} \left\{ \sum_{i=1}^N \lambda_i^2 - \sum_{j \neq k} \log |\lambda_j - \lambda_k| \right\} \right]$

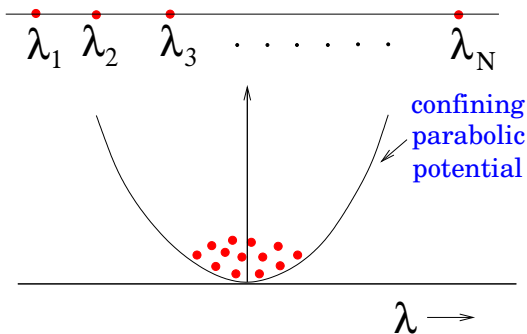
Coulomb Gas interpretation

- $Z_N = \int_{-\infty}^{\infty} \dots \int_{-\infty}^{\infty} \left\{ \prod_i d\lambda_i \right\} \exp \left[-\frac{\beta}{2} \left\{ \sum_{i=1}^N \lambda_i^2 - \sum_{j \neq k} \log |\lambda_j - \lambda_k| \right\} \right]$
- 2-d Coulomb gas confined to a line (Dyson) with $\beta \rightarrow$ inverse temp.



Coulomb Gas interpretation

- $Z_N = \int_{-\infty}^{\infty} \dots \int_{-\infty}^{\infty} \left\{ \prod_i d\lambda_i \right\} \exp \left[-\frac{\beta}{2} \left\{ \sum_{i=1}^N \lambda_i^2 - \sum_{j \neq k} \log |\lambda_j - \lambda_k| \right\} \right]$
- 2-d Coulomb gas confined to a line (Dyson) with $\beta \rightarrow$ inverse temp.



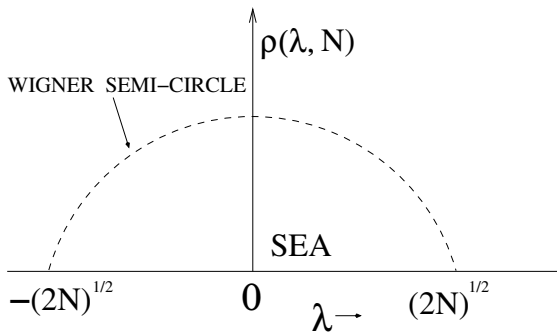
- Balance of energy $\rightarrow N \lambda^2 \sim N^2$
- Typical eigenvalue: $\lambda_{\text{typ}} \sim \sqrt{N}$ for large N

Spectral Density: Wigner's Semicircle Law

- Av. density of states: $\rho(\lambda, N) = \left\langle \frac{1}{N} \sum_{i=1}^N \delta(\lambda - \lambda_i) \right\rangle$

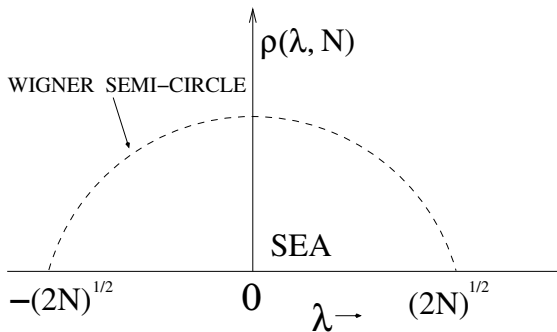
Spectral Density: Wigner's Semicircle Law

- Av. density of states: $\rho(\lambda, N) = \langle \frac{1}{N} \sum_{i=1}^N \delta(\lambda - \lambda_i) \rangle$
- Wigner's Semi-circle: $\rho(\lambda, N) \xrightarrow{N \rightarrow \infty} \sqrt{\frac{2}{N\pi^2}} \left[1 - \frac{\lambda^2}{2N} \right]^{1/2}$



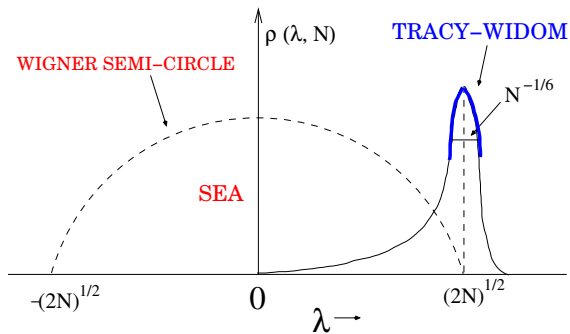
Spectral Density: Wigner's Semicircle Law

- Av. density of states: $\rho(\lambda, N) = \left\langle \frac{1}{N} \sum_{i=1}^N \delta(\lambda - \lambda_i) \right\rangle$
- Wigner's Semi-circle: $\rho(\lambda, N) \xrightarrow{N \rightarrow \infty} \sqrt{\frac{2}{N\pi^2}} \left[1 - \frac{\lambda^2}{2N} \right]^{1/2}$

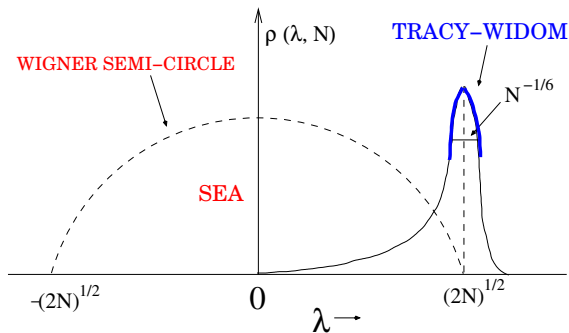


- $\langle \lambda_{\max} \rangle = \sqrt{2N}$ for large N .
- λ_{\max} fluctuates from one sample to another. $\text{Prob}[\lambda_{\max}, N] = ?$

Tracy-Widom distribution for λ_{\max}

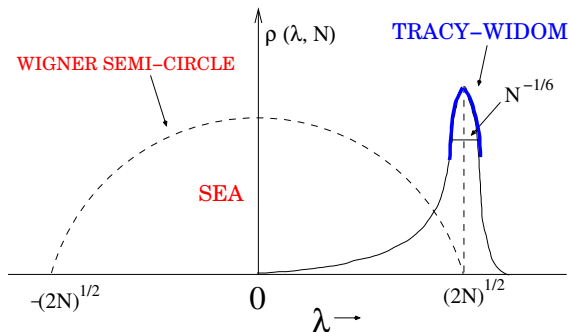


Tracy-Widom distribution for λ_{\max}



- $\langle \lambda_{\max} \rangle = \sqrt{2N}$; typical fluctuation: $|\lambda_{\max} - \sqrt{2N}| \sim N^{-1/6}$ (small)

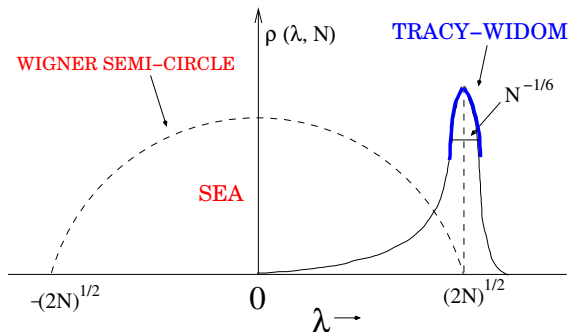
Tracy-Widom distribution for λ_{\max}



- $\langle \lambda_{\max} \rangle = \sqrt{2N}$; typical fluctuation: $|\lambda_{\max} - \sqrt{2N}| \sim N^{-1/6}$ (small)
- typical fluctuations are distributed via Tracy-Widom (1994):
- cumulative distribution:

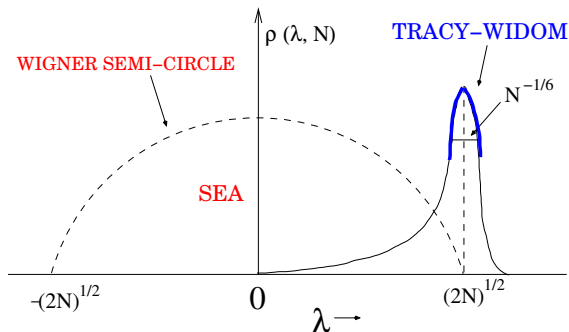
$$\text{Prob}[\lambda_{\max} \leq t, N] \rightarrow F_{\beta} \left(\sqrt{2}N^{1/6}(t - \sqrt{2N}) \right)$$

Tracy-Widom distribution for λ_{\max}



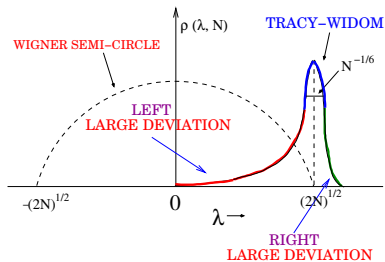
- $\langle \lambda_{\max} \rangle = \sqrt{2N}$; typical fluctuation: $|\lambda_{\max} - \sqrt{2N}| \sim N^{-1/6}$ (small)
- typical fluctuations are distributed via Tracy-Widom (1994):
- cumulative distribution:
$$\text{Prob}[\lambda_{\max} \leq t, N] \rightarrow F_{\beta} \left(\sqrt{2}N^{1/6}(t - \sqrt{2N}) \right)$$
- Prob. density (pdf): $f_{\beta}(z) = dF_{\beta}(z)/dz$

Tracy-Widom distribution for λ_{\max}

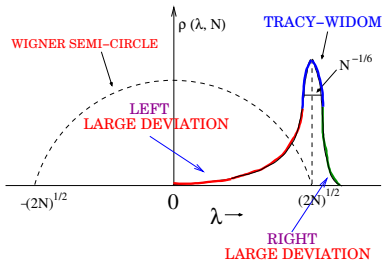


- $\langle \lambda_{\max} \rangle = \sqrt{2N}$; typical fluctuation: $|\lambda_{\max} - \sqrt{2N}| \sim N^{-1/6}$ (small)
- typical fluctuations are distributed via Tracy-Widom (1994):
- cumulative distribution:
$$\text{Prob}[\lambda_{\max} \leq t, N] \rightarrow F_{\beta} \left(\sqrt{2}N^{1/6}(t - \sqrt{2N}) \right)$$
- Prob. density (pdf): $f_{\beta}(z) = dF_{\beta}(z)/dz$
- $F_{\beta}(z) \rightarrow$ solution of Painlevé equation

Probability of Large Deviations of λ_{\max} :

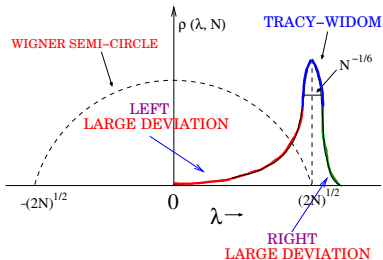


Probability of Large Deviations of λ_{\max} :



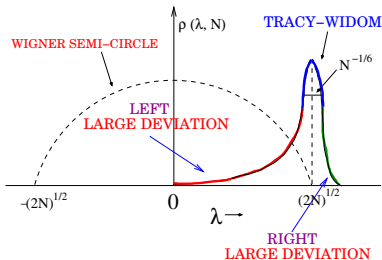
- Tracy-Widom law $\text{Prob}[\lambda_{\max} \leq t, N] \rightarrow F_{\beta} \left(\sqrt{2}N^{1/6}(t - \sqrt{2N}) \right)$ describes the prob. of **typical (small)** fluctuations of $\sim O(N^{-1/6})$ around the mean $\sqrt{2N}$, i.e., when $|\lambda_{\max} - \sqrt{2N}| \sim N^{-1/6}$

Probability of Large Deviations of λ_{\max} :



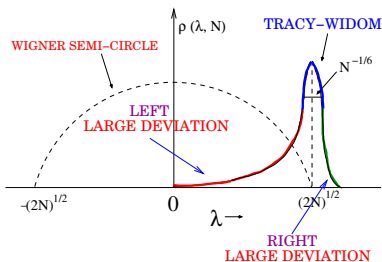
- Tracy-Widom law $\text{Prob}[\lambda_{\max} \leq t, N] \rightarrow F_{\beta}(\sqrt{2N}^{1/6}(t - \sqrt{2N}))$ describes the prob. of **typical (small)** fluctuations of $\sim O(N^{-1/6})$ around the mean $\sqrt{2N}$, i.e., when $|\lambda_{\max} - \sqrt{2N}| \sim N^{-1/6}$
- **Q**: How to describe the prob. of **large (atypical)** fluctuations?

Probability of Large Deviations of λ_{\max} :



- Tracy-Widom law $\text{Prob}[\lambda_{\max} \leq t, N] \rightarrow F_{\beta}(\sqrt{2N}^{1/6}(t - \sqrt{2N}))$ describes the prob. of **typical (small)** fluctuations of $\sim O(N^{-1/6})$ around the mean $\sqrt{2N}$, i.e., when $|\lambda_{\max} - \sqrt{2N}| \sim N^{-1/6}$
- **Q:** How to describe the prob. of **large (atypical)** fluctuations?
- Ex: $P_N = \text{Prob}[\lambda_{\max} \leq 0, N] = \text{Prob}[\lambda_1 \leq 0, \lambda_2 \leq 0, \dots, \lambda_N \leq 0]$

Probability of Large Deviations of λ_{\max} :



- Tracy-Widom law $\text{Prob}[\lambda_{\max} \leq t, N] \rightarrow F_\beta \left(\sqrt{2}N^{1/6}(t - \sqrt{2N}) \right)$ describes the prob. of **typical (small)** fluctuations of $\sim O(N^{-1/6})$ around the mean $\sqrt{2N}$, i.e., when $|\lambda_{\max} - \sqrt{2N}| \sim N^{-1/6}$
- **Q:** How to describe the prob. of **large (atypical)** fluctuations?
- Ex: $P_N = \text{Prob}[\lambda_{\max} \leq 0, N] = \text{Prob}[\lambda_1 \leq 0, \lambda_2 \leq 0, \dots, \lambda_N \leq 0]$
 $\rightarrow t = 0$, i.e., fluctuation = $t - \sqrt{2N} = -\sqrt{2N} \rightarrow$ **large**

Exact Left Large Deviation Function

- For large $\sim O(\sqrt{N})$ negative deviation: $\sqrt{2N} - t \sim O(\sqrt{N})$

$$\text{Prob}[\lambda_{\max} \leq t, N] \sim \exp \left[-\beta N^2 \phi_- \left(\frac{\sqrt{2N} - t}{\sqrt{N}} \right) \right]$$

Exact Left Large Deviation Function

- For large $\sim O(\sqrt{N})$ **negative** deviation: $\sqrt{2N} - t \sim O(\sqrt{N})$

$$\text{Prob}[\lambda_{\max} \leq t, N] \sim \exp \left[-\beta N^2 \Phi_- \left(\frac{\sqrt{2N} - t}{\sqrt{N}} \right) \right]$$

- where $\Phi_-(y) \rightarrow$ **left** large deviation function:

$$\begin{aligned} \Phi_-(y = z + \sqrt{2}) &= \frac{1}{108} \left[36z^2 - z^4 + (15z + z^3)\sqrt{z^2 + 6} \right. \\ &\quad \left. + 27 \left(\ln(18) - 2 \ln(-z + \sqrt{6 + z^2}) \right) \right] \end{aligned}$$

(D.Dean & S.M., 2006, 2008)

Exact Left Large Deviation Function

- For large $\sim O(\sqrt{N})$ **negative** deviation: $\sqrt{2N} - t \sim O(\sqrt{N})$

$$\text{Prob}[\lambda_{\max} \leq t, N] \sim \exp \left[-\beta N^2 \Phi_- \left(\frac{\sqrt{2N} - t}{\sqrt{N}} \right) \right]$$

- where $\Phi_-(y) \rightarrow$ **left** large deviation function:

$$\begin{aligned} \Phi_-(y = z + \sqrt{2}) &= \frac{1}{108} \left[36z^2 - z^4 + (15z + z^3)\sqrt{z^2 + 6} \right. \\ &\quad \left. + 27 \left(\ln(18) - 2 \ln(-z + \sqrt{6 + z^2}) \right) \right] \end{aligned}$$

(D. Dean & S.M., 2006, 2008)

- In particular, with $t = 0$,

$$P_N = \text{Prob}[\lambda_{\max} \leq 0, N] \sim \exp[-\beta \theta N^2]$$

$$\theta = \Phi_-(\sqrt{2}) = \frac{1}{4} \ln(3) = 0.274653 \dots$$

Exact Right Large Deviation Function

- For large $\sim O(\sqrt{N})$ positive deviation: $t - \sqrt{2N} \sim O(\sqrt{N})$

$$\text{Prob}[\lambda_{\max} = t, N] \sim \exp \left[-\beta N \Phi_+ \left(\frac{t - \sqrt{2N}}{\sqrt{N}} \right) \right]$$

Exact Right Large Deviation Function

- For large $\sim O(\sqrt{N})$ positive deviation: $t - \sqrt{2N} \sim O(\sqrt{N})$

$$\text{Prob}[\lambda_{\max} = t, N] \sim \exp \left[-\beta N \Phi_+ \left(\frac{t - \sqrt{2N}}{\sqrt{N}} \right) \right]$$

- where $\Phi_+(y) \rightarrow$ right large deviation function:

$$\Phi_+(y = z - \sqrt{2}) = \frac{z^2 - 1}{2} - \ln(z\sqrt{2}) + \frac{1}{4z^2} G(2/z^2)$$

where $G(z) = {}_3F_2[\{1, 1, 3/2\}, 2, 3, z] \rightarrow$ Hypergeometric function

(S.M. & Vergassola, 2009)

Exact Right Large Deviation Function

- For large $\sim O(\sqrt{N})$ positive deviation: $t - \sqrt{2N} \sim O(\sqrt{N})$

$$\text{Prob}[\lambda_{\max} = t, N] \sim \exp \left[-\beta N \Phi_+ \left(\frac{t - \sqrt{2N}}{\sqrt{N}} \right) \right]$$

- where $\Phi_+(y) \rightarrow$ right large deviation function:

$$\Phi_+(y = z - \sqrt{2}) = \frac{z^2 - 1}{2} - \ln(z\sqrt{2}) + \frac{1}{4z^2} G(2/z^2)$$

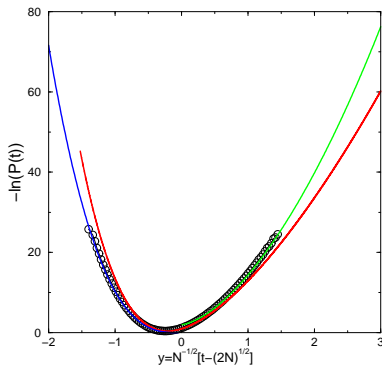
where $G(z) = {}_3F_2[\{1, 1, 3/2\}, 2, 3, z] \rightarrow$ Hypergeometric function

(S.M. & Vergassola, 2009)

- In particular,

$$\begin{aligned} \Phi_+(y) &\approx \frac{2^{7/4}}{3} y^{3/2} \quad \text{as } y \rightarrow 0 \\ &\approx y^2/2 \quad \text{as } y \rightarrow \infty \end{aligned}$$

Comparison with Simulations:



$N \times N$ real Gaussian matrix ($\beta = 1$): $N = 10$ (S.M. & Vergassola, 2009)

circles \rightarrow simulation points

red line \rightarrow Tracy-Widom

blue line \rightarrow left large deviation function ($\propto N^2$)

green line \rightarrow right large deviation function ($\propto N$).

Left Large Deviation: Sketch of the Method

$$\text{Prob}[\lambda_{\max} \leq t, N] = \frac{Z_N(t)}{Z_N(\infty)}$$

$$Z_N(t) = \int_{-\infty}^t \cdots \int_{-\infty}^t \left\{ \prod_i d\lambda_i \right\} \exp \left[-\frac{\beta}{2} \left\{ \sum_{i=1}^N \lambda_i^2 - \sum_{j \neq k} \log |\lambda_j - \lambda_k| \right\} \right]$$

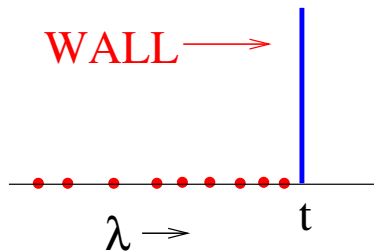
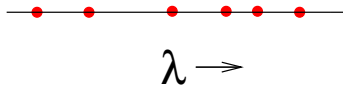
Left Large Deviation: Sketch of the Method

$$\text{Prob}[\lambda_{\max} \leq t, N] = \frac{Z_N(t)}{Z_N(\infty)}$$

$$Z_N(t) = \int_{-\infty}^t \cdots \int_{-\infty}^t \left\{ \prod_i d\lambda_i \right\} \exp \left[-\frac{\beta}{2} \left\{ \sum_{i=1}^N \lambda_i^2 - \sum_{j \neq k} \log |\lambda_j - \lambda_k| \right\} \right]$$

denominator

numerator



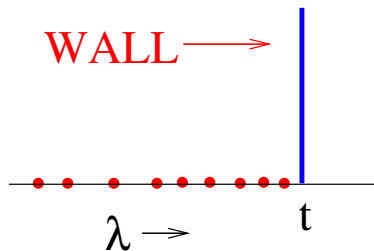
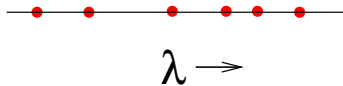
Left Large Deviation: Sketch of the Method

$$\text{Prob}[\lambda_{\max} \leq t, N] = \frac{Z_N(t)}{Z_N(\infty)}$$

$$Z_N(t) = \int_{-\infty}^t \cdots \int_{-\infty}^t \left\{ \prod_i d\lambda_i \right\} \exp \left[-\frac{\beta}{2} \left\{ \sum_{i=1}^N \lambda_i^2 - \sum_{j \neq k} \log |\lambda_j - \lambda_k| \right\} \right]$$

denominator

numerator



Multidimensional Integral \rightarrow Functional Integral

- Scaled variables: $x_i = \lambda_i / \sqrt{N}$; maximum x_i : $w = t / \sqrt{N}$

Multidimensional Integral \rightarrow Functional Integral

- Scaled variables: $x_i = \lambda_i / \sqrt{N}$; maximum x_i : $w = t / \sqrt{N}$

- $$Z_N(t) \propto \int_{-\infty}^t \prod_i dx_i \exp [-\beta N^2 H(\{x_i\})]$$

$$H(\{x_i\}) = \frac{1}{2N} \sum_i x_i^2 - \frac{1}{2N^2} \sum_{j \neq k} \log |x_j - x_k|$$

Multidimensional Integral \rightarrow Functional Integral

- Scaled variables: $x_i = \lambda_i / \sqrt{N}$; maximum x_j : $w = t / \sqrt{N}$

- $$Z_N(t) \propto \int_{-\infty}^t \prod_i dx_i \exp [-\beta N^2 H(\{x_i\})]$$

$$H(\{x_i\}) = \frac{1}{2N} \sum_i x_i^2 - \frac{1}{2N^2} \sum_{j \neq k} \log |x_j - x_k|$$

- Introduce counting function (scaled density): $f(x) = \frac{1}{N} \sum_i \delta(x - x_i)$

Multidimensional Integral \rightarrow Functional Integral

- Scaled variables: $x_i = \lambda_i / \sqrt{N}$; maximum x_j : $w = t / \sqrt{N}$

- $$Z_N(t) \propto \int_{-\infty}^t \prod_i dx_i \exp[-\beta N^2 H(\{x_i\})]$$

$$H(\{x_i\}) = \frac{1}{2N} \sum_i x_i^2 - \frac{1}{2N^2} \sum_{j \neq k} \log|x_j - x_k|$$

- Introduce counting function (scaled density): $f(x) = \frac{1}{N} \sum_i \delta(x - x_i)$
- discrete sum \rightarrow continuous integral:

$$H[f(x)] = \int_{-\infty}^w x^2 f(x) dx - \int_{-\infty}^w \int_{-\infty}^w \ln|x - x'| f(x) f(x') dx dx'$$

Multidimensional Integral \rightarrow Functional Integral

- Scaled variables: $x_i = \lambda_i / \sqrt{N}$; maximum x_i : $w = t / \sqrt{N}$

- $$Z_N(t) \propto \int_{-\infty}^t \prod_i dx_i \exp[-\beta N^2 H(\{x_i\})]$$

$$H(\{x_i\}) = \frac{1}{2N} \sum_i x_i^2 - \frac{1}{2N^2} \sum_{j \neq k} \log |x_j - x_k|$$

- Introduce counting function (scaled density): $f(x) = \frac{1}{N} \sum_i \delta(x - x_i)$
- discrete sum \rightarrow continuous integral:

$$H[f(x)] = \int_{-\infty}^w x^2 f(x) dx - \int_{-\infty}^w \int_{-\infty}^w \ln |x - x'| f(x) f(x') dx dx'$$

$$Z_N(w) \propto \int \mathcal{D}f(x) \exp \left[-\beta N^2 \left\{ H[f(x)] + C \left(\int f(x) dx - 1 \right) \right\} + O(N) \right]$$

Multidimensional Integral \rightarrow Functional Integral

- Scaled variables: $x_i = \lambda_i/\sqrt{N}$; maximum x_i : $w = t/\sqrt{N}$

- $$Z_N(t) \propto \int_{-\infty}^t \prod_i dx_i \exp[-\beta N^2 H(\{x_i\})]$$

$$H(\{x_i\}) = \frac{1}{2N} \sum_i x_i^2 - \frac{1}{2N^2} \sum_{j \neq k} \log|x_j - x_k|$$

- Introduce counting function (scaled density): $f(x) = \frac{1}{N} \sum_i \delta(x - x_i)$
- discrete sum \rightarrow continuous integral:

$$H[f(x)] = \int_{-\infty}^w x^2 f(x) dx - \int_{-\infty}^w \int_{-\infty}^w \ln|x - x'| f(x) f(x') dx dx'$$

$$Z_N(w) \propto \int \mathcal{D}f(x) \exp\left[-\beta N^2 \left\{ H[f(x)] + C \left(\int f(x) dx - 1 \right) \right\} + O(N)\right]$$

- for large N , minimize the action $S[f(x)] = H[f(x)] + C[\int f(x) dx - 1]$
(saddle point): $\frac{\delta S}{\delta f} = 0 \rightarrow f_w(x)$

Multidimensional Integral \rightarrow Functional Integral

- Scaled variables: $x_i = \lambda_i/\sqrt{N}$; maximum x_j : $w = t/\sqrt{N}$

- $$Z_N(t) \propto \int_{-\infty}^t \prod_i dx_i \exp[-\beta N^2 H(\{x_i\})]$$

$$H(\{x_i\}) = \frac{1}{2N} \sum_i x_i^2 - \frac{1}{2N^2} \sum_{j \neq k} \log|x_j - x_k|$$

- Introduce counting function (scaled density): $f(x) = \frac{1}{N} \sum_i \delta(x - x_i)$
- discrete sum \rightarrow continuous integral:

$$H[f(x)] = \int_{-\infty}^w x^2 f(x) dx - \int_{-\infty}^w \int_{-\infty}^w \ln|x - x'| f(x) f(x') dx dx'$$

$$Z_N(w) \propto \int \mathcal{D}f(x) \exp\left[-\beta N^2 \left\{ H[f(x)] + C \left(\int f(x) dx - 1 \right) \right\} + O(N)\right]$$

- for large N , minimize the action $S[f(x)] = H[f(x)] + C[\int f(x) dx - 1]$

(saddle point): $\frac{\delta S}{\delta f} = 0 \rightarrow f_w(x) \rightarrow$

$$Z_N(w) \sim \exp[-\beta N^2 S[f_w(x)]]$$

Saddle Point Solution

- saddle point $\frac{\delta S}{\delta f} = 0 \rightarrow$ singular integral Eq. for $f_w(x)$

Saddle Point Solution

- saddle point $\frac{\delta S}{\delta f} = 0 \rightarrow$ singular integral Eq. for $f_w(x)$

- $$x = \mathcal{P} \int_{-\infty}^w \frac{f_w(y) dy}{x - y} \rightarrow \text{Semi-Hilbert transform}$$

\rightarrow Inverse electrostatic problem \rightarrow Given the potential x find the charge density $f_w(x)$

Saddle Point Solution

- saddle point $\frac{\delta S}{\delta f} = 0 \rightarrow$ singular integral Eq. for $f_w(x)$

- $$x = \mathcal{P} \int_{-\infty}^w \frac{f_w(y) dy}{x - y} \rightarrow \text{Semi-Hilbert transform}$$

\rightarrow Inverse electrostatic problem \rightarrow Given the potential x find the charge density $f_w(x)$

- General method for solving such singular integral equations \rightarrow Tricomi (1957)

Exact Saddle Point Solution

- Exact solution (D. Dean and S.M., 2008):

$$f_w(x) = \frac{\sqrt{x + L_1(w)}}{2\pi\sqrt{w - x}} [w + L_1(w) - 2x]$$

where $-L_1(w) \leq x \leq w$ and $L_1(w) = [2\sqrt{w^2 + 6} - w]/3$

Exact Saddle Point Solution

- Exact solution (D. Dean and S.M., 2008):

$$f_w(x) = \frac{\sqrt{x + L_1(w)}}{2\pi\sqrt{w - x}} [w + L_1(w) - 2x]$$

where $-L_1(w) \leq x \leq w$ and $L_1(w) = [2\sqrt{w^2 + 6} - w]/3$

- When $w \rightarrow \infty$, $L_1(w) \rightarrow \sqrt{2}$ and $f_\infty(x) = \sqrt{2 - x^2}/\pi \rightarrow$ semicircle

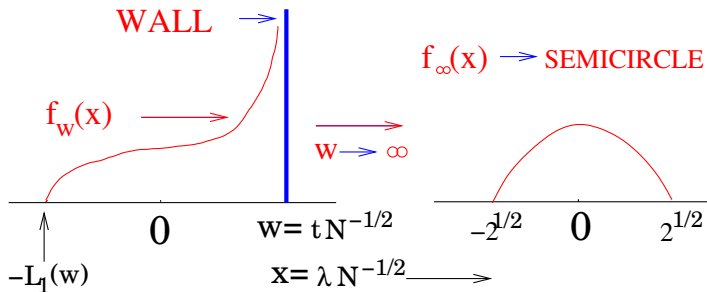
Exact Saddle Point Solution

- Exact solution (D. Dean and S.M., 2008):

$$f_w(x) = \frac{\sqrt{x + L_1(w)}}{2\pi\sqrt{w-x}} [w + L_1(w) - 2x]$$

where $-L_1(w) \leq x \leq w$ and $L_1(w) = [2\sqrt{w^2 + 6} - w]/3$

- When $w \rightarrow \infty$, $L_1(w) \rightarrow \sqrt{2}$ and $f_\infty(x) = \sqrt{2-x^2}/\pi \rightarrow$ semicircle



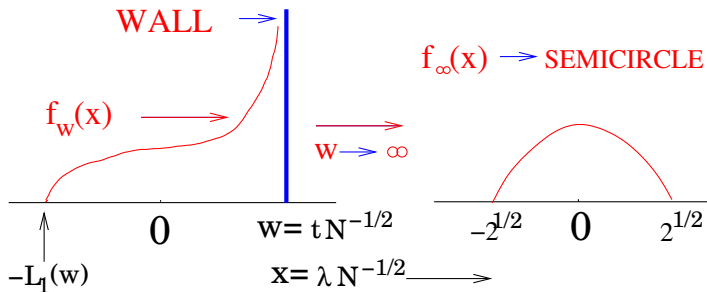
Exact Saddle Point Solution

- Exact solution (D. Dean and S.M., 2008):

$$f_w(x) = \frac{\sqrt{x + L_1(w)}}{2\pi\sqrt{w-x}} [w + L_1(w) - 2x]$$

where $-L_1(w) \leq x \leq w$ and $L_1(w) = [2\sqrt{w^2 + 6} - w]/3$

- When $w \rightarrow \infty$, $L_1(w) \rightarrow \sqrt{2}$ and $f_\infty(x) = \sqrt{2-x^2}/\pi \rightarrow$ semicircle



Large Deviation Function

- $$\begin{aligned}\text{Prob}[\lambda_{\max} \leq t, N] &= \frac{Z_N(t)}{Z_N(\infty)} \\ &\sim \exp \left[-\beta N^2 \left\{ S[f_{w=t/\sqrt{N}}(x)] - S[f_{\infty}(x)] \right\} \right] \\ &\sim \exp \left[-\beta N^2 \Phi_- \left(\frac{\sqrt{2N} - t}{\sqrt{N}} \right) \right]\end{aligned}$$

Large Deviation Function

- $$\begin{aligned}\text{Prob}[\lambda_{\max} \leq t, N] &= \frac{Z_N(t)}{Z_N(\infty)} \\ &\sim \exp \left[-\beta N^2 \left\{ S[f_{w=t/\sqrt{N}}(x)] - S[f_{\infty}(x)] \right\} \right] \\ &\sim \exp \left[-\beta N^2 \Phi_- \left(\frac{\sqrt{2N} - t}{\sqrt{N}} \right) \right]\end{aligned}$$

• where $\Phi_-(y) \rightarrow$ left large deviation function:

$$\begin{aligned}\Phi_-(y = z + \sqrt{2}) &= \frac{1}{108} \left[36z^2 - z^4 + (15z + z^3)\sqrt{z^2 + 6} \right. \\ &\quad \left. + 27 \left(\ln(18) - 2 \ln(-z + \sqrt{6 + z^2}) \right) \right]\end{aligned}$$

Large Deviation Function

- $$\begin{aligned}\text{Prob}[\lambda_{\max} \leq t, N] &= \frac{Z_N(t)}{Z_N(\infty)} \\ &\sim \exp \left[-\beta N^2 \left\{ S[f_{w=t/\sqrt{N}}(x)] - S[f_{\infty}(x)] \right\} \right] \\ &\sim \exp \left[-\beta N^2 \Phi_- \left(\frac{\sqrt{2N} - t}{\sqrt{N}} \right) \right]\end{aligned}$$

- where $\Phi_-(y) \rightarrow$ left large deviation function:

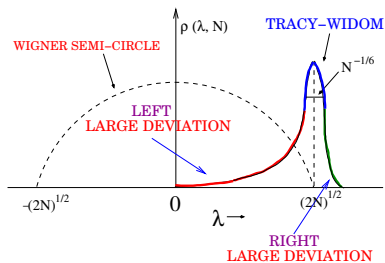
$$\begin{aligned}\Phi_-(y = z + \sqrt{2}) &= \frac{1}{108} \left[36z^2 - z^4 + (15z + z^3)\sqrt{z^2 + 6} \right. \\ &\quad \left. + 27 \left(\ln(18) - 2 \ln(-z + \sqrt{6 + z^2}) \right) \right]\end{aligned}$$

- In particular, for $t = 0$, $P_N \sim \exp[-\beta\theta N^2]$

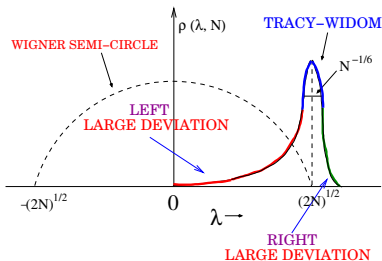
$$\theta = \Phi_-(\sqrt{2}) = \frac{1}{4} \ln(3) = 0.274653 \dots$$

(D. Dean & S.M., 2006, 2008)

Matching the left tail of Tracy-Widom distribution:

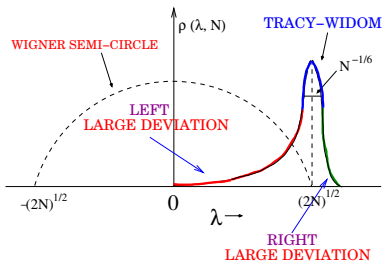


Matching the left tail of Tracy-Widom distribution:



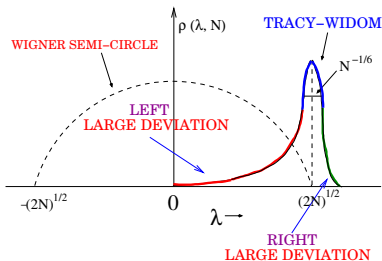
- $\text{Prob}[\lambda_{\max} \leq t, N] \sim \exp\left[-\beta N^2 \Phi_{-}\left(\frac{\sqrt{2N}-t}{\sqrt{N}}\right)\right]; \quad y = \frac{\sqrt{2N}-t}{\sqrt{N}}$

Matching the left tail of Tracy-Widom distribution:



- $\text{Prob}[\lambda_{\max} \leq t, N] \sim \exp \left[-\beta N^2 \Phi_- \left(\frac{\sqrt{2N}-t}{\sqrt{N}} \right) \right]; \quad y = \frac{\sqrt{2N}-t}{\sqrt{N}}$
- When $y \ll 1$, i.e., $\sqrt{2N}-t \ll \sqrt{N} \rightarrow$ left tail of **Tracy-Widom**

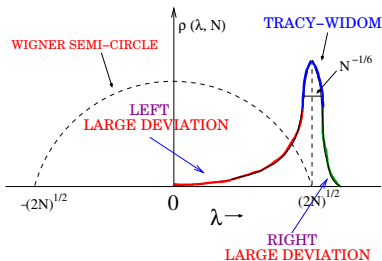
Matching the left tail of Tracy-Widom distribution:



- $\text{Prob}[\lambda_{\max} \leq t, N] \sim \exp \left[-\beta N^2 \Phi_- \left(\frac{\sqrt{2N}-t}{\sqrt{N}} \right) \right]; \quad y = \frac{\sqrt{2N}-t}{\sqrt{N}}$
- When $y \ll 1$, i.e., $\sqrt{2N}-t \ll \sqrt{N} \rightarrow$ left tail of **Tracy-Widom**
- As $y \rightarrow 0$, $\Phi_-(y) \rightarrow \frac{y^3}{6\sqrt{2}} \Rightarrow$

$$\text{Prob}[\lambda_{\max} \leq t, N] \approx \exp \left[-\frac{\beta}{24} \left| \sqrt{2} N^{1/6} (t - \sqrt{2N}) \right|^3 \right]$$

Matching the left tail of Tracy-Widom distribution:

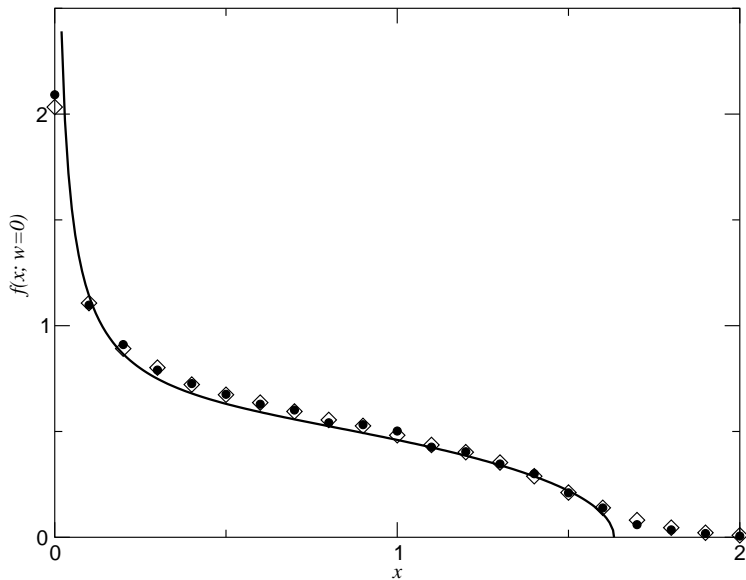


- $\text{Prob}[\lambda_{\max} \leq t, N] \sim \exp \left[-\beta N^2 \Phi_- \left(\frac{\sqrt{2N}-t}{\sqrt{N}} \right) \right]; \quad y = \frac{\sqrt{2N}-t}{\sqrt{N}}$
- When $y \ll 1$, i.e., $\sqrt{2N}-t \ll \sqrt{N} \rightarrow$ left tail of **Tracy-Widom**
- As $y \rightarrow 0$, $\Phi_-(y) \rightarrow \frac{y^3}{6\sqrt{2}} \Rightarrow$

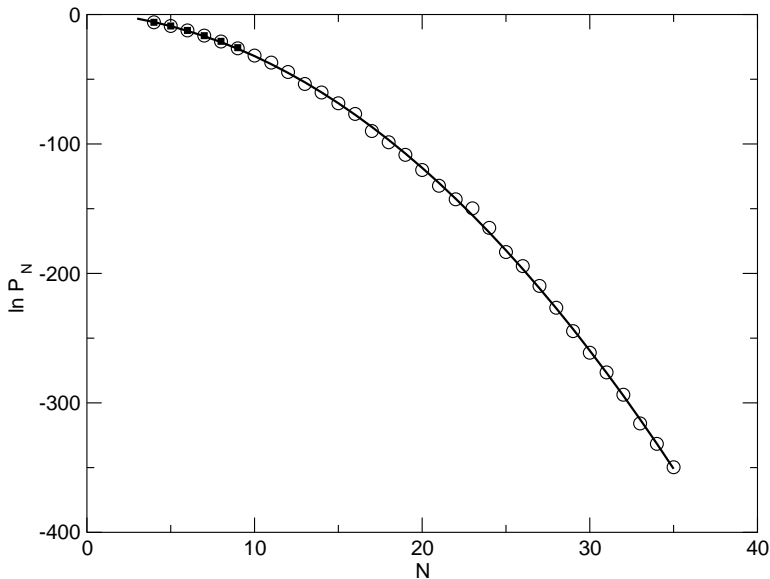
$$\text{Prob}[\lambda_{\max} \leq t, N] \approx \exp \left[-\frac{\beta}{24} \left| \sqrt{2} N^{1/6} (t - \sqrt{2N}) \right|^3 \right]$$

- recovers the correct left tail of **TW**: $F_\beta(x) \sim \exp[-\frac{\beta}{24} |x|^3]$ as $x \rightarrow -\infty$

Numerical Results: Charge Density at $w = 0$



Numerical Result for P_N :



Right Large Deviation Function

- For large $\sim O(\sqrt{N})$ positive deviation: $t - \sqrt{2N} \sim O(\sqrt{N})$

$$\text{Prob}[\lambda_{\max} = t, N] \sim \exp \left[-\beta N \Phi_+ \left(\frac{t - \sqrt{2N}}{\sqrt{N}} \right) \right]$$

Right Large Deviation Function

- For large $\sim O(\sqrt{N})$ positive deviation: $t - \sqrt{2N} \sim O(\sqrt{N})$

$$\text{Prob}[\lambda_{\max} = t, N] \sim \exp \left[-\beta N \Phi_+ \left(\frac{t - \sqrt{2N}}{\sqrt{N}} \right) \right]$$

- $P(\lambda_{\max} = t, N) \propto \exp[-\beta \Delta E(t)/2]$ where

Right Large Deviation Function

- For large $\sim O(\sqrt{N})$ positive deviation: $t - \sqrt{2N} \sim O(\sqrt{N})$

$$\text{Prob}[\lambda_{\max} = t, N] \sim \exp \left[-\beta N \Phi_+ \left(\frac{t - \sqrt{2N}}{\sqrt{N}} \right) \right]$$

- $P(\lambda_{\max} = t, N) \propto \exp[-\beta \Delta E(t)/2]$ where

$$\Delta E(t) = t^2 - 2N \int_{-\sqrt{2N}}^{\sqrt{2N}} \ln(t - \lambda) \rho_N(\lambda) d\lambda$$

\Rightarrow energy cost in pulling a charge out of the Wigner sea

Right Large Deviation Function

- For large $\sim O(\sqrt{N})$ positive deviation: $t - \sqrt{2N} \sim O(\sqrt{N})$

$$\text{Prob}[\lambda_{\max} = t, N] \sim \exp \left[-\beta N \Phi_+ \left(\frac{t - \sqrt{2N}}{\sqrt{N}} \right) \right]$$

- $P(\lambda_{\max} = t, N) \propto \exp[-\beta \Delta E(t)/2]$ where

$$\Delta E(t) = t^2 - 2N \int_{-\sqrt{2N}}^{\sqrt{2N}} \ln(t - \lambda) \rho_N(\lambda) d\lambda$$

\implies energy cost in pulling a charge out of the Wigner sea

$$\rho(\lambda, N) \xrightarrow{N \rightarrow \infty} \sqrt{\frac{2}{N\pi^2}} \left[1 - \frac{\lambda^2}{2N} \right]^{1/2} \rightarrow \text{av. den.}$$

Right Large Deviation Function

- For large $\sim O(\sqrt{N})$ positive deviation: $t - \sqrt{2N} \sim O(\sqrt{N})$

$$\text{Prob}[\lambda_{\max} = t, N] \sim \exp \left[-\beta N \Phi_+ \left(\frac{t - \sqrt{2N}}{\sqrt{N}} \right) \right]$$

- $P(\lambda_{\max} = t, N) \propto \exp[-\beta \Delta E(t)/2]$ where

$$\Delta E(t) = t^2 - 2N \int_{-\sqrt{2N}}^{\sqrt{2N}} \ln(t - \lambda) \rho_N(\lambda) d\lambda$$

\Rightarrow energy cost in pulling a charge out of the Wigner sea

$$\rho(\lambda, N) \xrightarrow{N \rightarrow \infty} \sqrt{\frac{2}{N\pi^2}} \left[1 - \frac{\lambda^2}{2N} \right]^{1/2} \rightarrow \text{av. den.}$$

\Rightarrow exact $\Phi_+(y)$ (S.M. & Vergassola, 2009)

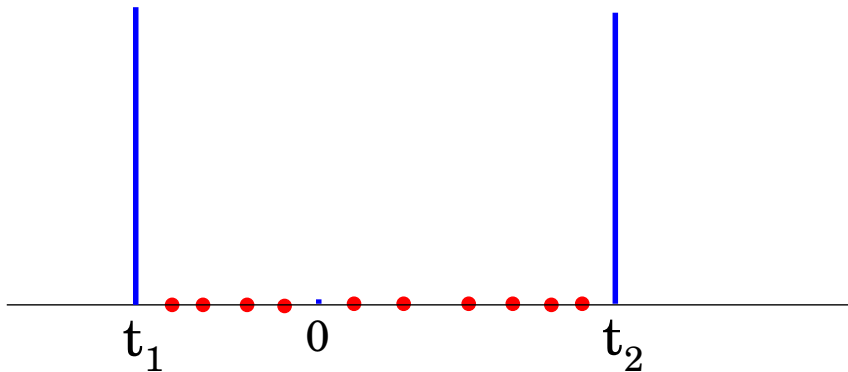
For $\beta = 1$, $\Phi_+(y) \rightarrow$ different method by Ben-Arous et. al. (2001).

Joint distribution of λ_{\min} and λ_{\max} :

- $\text{Prob}[\lambda_{\min} \geq t_1, \lambda_{\max} \leq t_2] = \frac{Z_N(t_1, t_2)}{Z_N(-\infty, \infty)}$
- $Z_N(t_1, t_2) =$ Partition function with *two* walls

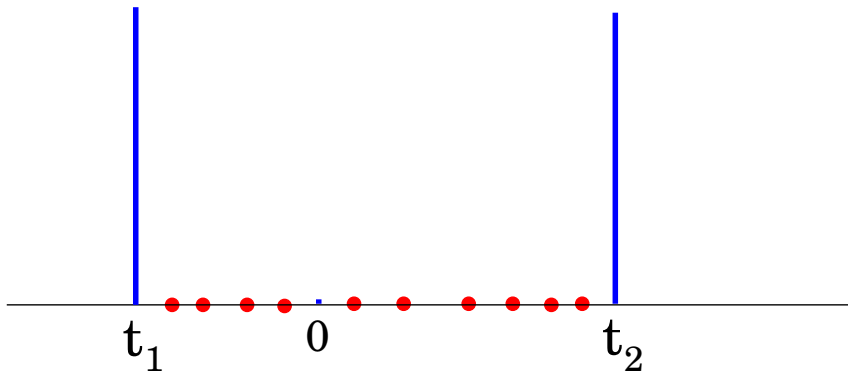
Joint distribution of λ_{\min} and λ_{\max} :

- $\text{Prob}[\lambda_{\min} \geq t_1, \lambda_{\max} \leq t_2] = \frac{Z_N(t_1, t_2)}{Z_N(-\infty, \infty)}$
- $Z_N(t_1, t_2) =$ Partition function with *two* walls



Joint distribution of λ_{\min} and λ_{\max} :

- $\text{Prob}[\lambda_{\min} \geq t_1, \lambda_{\max} \leq t_2] = \frac{Z_N(t_1, t_2)}{Z_N(-\infty, \infty)}$
- $Z_N(t_1, t_2) =$ Partition function with *two* walls



Joint distribution of λ_{\min} and λ_{\max} :

- When both $t_1 \sim O(\sqrt{N})$ and $t_2 \sim O(\sqrt{N})$

Joint distribution of λ_{\min} and λ_{\max} :

- When both $t_1 \sim O(\sqrt{N})$ and $t_2 \sim O(\sqrt{N})$

$$\text{Prob}[\lambda_{\min} \geq t_1, \lambda_{\max} \leq t_2] = \exp \left[-\beta N^2 \Psi \left(\frac{t_1}{\sqrt{N}}, \frac{t_2}{\sqrt{N}} \right) + O(N) \right]$$

Joint distribution of λ_{\min} and λ_{\max} :

- When both $t_1 \sim O(\sqrt{N})$ and $t_2 \sim O(\sqrt{N})$

$$\text{Prob}[\lambda_{\min} \geq t_1, \lambda_{\max} \leq t_2] = \exp \left[-\beta N^2 \Psi \left(\frac{t_1}{\sqrt{N}}, \frac{t_2}{\sqrt{N}} \right) + O(N) \right]$$

Joint distribution of λ_{\min} and λ_{\max} :

- When both $t_1 \sim O(\sqrt{N})$ and $t_2 \sim O(\sqrt{N})$

$$\text{Prob}[\lambda_{\min} \geq t_1, \lambda_{\max} \leq t_2] = \exp \left[-\beta N^2 \Psi \left(\frac{t_1}{\sqrt{N}}, \frac{t_2}{\sqrt{N}} \right) + O(N) \right]$$

- $\Psi(z_1, z_2) = U(z_1, z_2 - z_1) - \frac{3+\ln(4)}{8}$

$$U(x, y) = \frac{1}{32} \left[32 \ln(2) - 16 \ln(y) + 16x^2 + 6y^2 \right. \\ \left. + 16yx - 2x^2y^2 - 2y^3x - \frac{9}{16}y^4 \right]$$

(D.Dean & S.M., 2008)

Largest Eigenvalue of Wishart Matrix

- Let $X = [x_{i,j}] \rightarrow (M \times N)$ rectangular matrix

Largest Eigenvalue of Wishart Matrix

- Let $X = [x_{i,j}] \rightarrow (M \times N)$ rectangular matrix
- $W = X^\dagger X \rightarrow (N \times N)$ square **covariance** matrix (**Wishart**)

Largest Eigenvalue of Wishart Matrix

- Let $X = [x_{i,j}] \rightarrow (M \times N)$ rectangular matrix
- $W = X^\dagger X \rightarrow (N \times N)$ square **covariance** matrix (**Wishart**)
- Entries of X Gaussian: $\Pr[X] \propto \exp \left[-\frac{\beta}{2} \text{Tr}(X^\dagger X) \right]$

Largest Eigenvalue of Wishart Matrix

- Let $X = [x_{i,j}] \rightarrow (M \times N)$ rectangular matrix
- $W = X^\dagger X \rightarrow (N \times N)$ square **covariance** matrix (**Wishart**)
- Entries of X Gaussian: $\Pr[X] \propto \exp \left[-\frac{\beta}{2} \text{Tr}(X^\dagger X) \right]$
- N real eigenvalues of W : $\lambda_1 \geq 0, \lambda_2 \geq 0, \dots, \lambda_N \geq 0$

Largest Eigenvalue of Wishart Matrix

- Let $X = [x_{i,j}] \rightarrow (M \times N)$ rectangular matrix
- $W = X^\dagger X \rightarrow (N \times N)$ square **covariance** matrix (**Wishart**)
- Entries of X Gaussian: $\Pr[X] \propto \exp \left[-\frac{\beta}{2} \text{Tr}(X^\dagger X) \right]$
- N real eigenvalues of W : $\lambda_1 \geq 0, \lambda_2 \geq 0, \dots, \lambda_N \geq 0$
- Joint distribution of eigenvalues (James, 1964):

$$P(\{\lambda_i\}) \propto \exp \left[-\frac{\beta}{2} \sum_{i=1}^N \lambda_i \right] \prod_i \lambda_i^{\frac{\beta}{2}(1+M-N)-1} \prod_{j < k} |\lambda_j - \lambda_k|^\beta$$

Largest Eigenvalue of Wishart Matrix

- Let $X = [x_{i,j}] \rightarrow (M \times N)$ rectangular matrix
- $W = X^\dagger X \rightarrow (N \times N)$ square **covariance** matrix (**Wishart**)
- Entries of X Gaussian: $\Pr[X] \propto \exp \left[-\frac{\beta}{2} \text{Tr}(X^\dagger X) \right]$
- N real eigenvalues of W : $\lambda_1 \geq 0, \lambda_2 \geq 0, \dots, \lambda_N \geq 0$
- Joint distribution of eigenvalues (James, 1964):

$$P(\{\lambda_i\}) \propto \exp \left[-\frac{\beta}{2} \sum_{i=1}^N \lambda_i \right] \prod_i \lambda_i^{\frac{\beta}{2}(1+M-N)-1} \prod_{j < k} |\lambda_j - \lambda_k|^\beta$$

- Av. density of states: $\rho(\lambda, N) = \left\langle \frac{1}{N} \sum_{i=1}^N \delta(\lambda - \lambda_i) \right\rangle \xrightarrow{N \rightarrow \infty} \frac{1}{N} f_{\text{MP}}\left(\frac{\lambda}{N}\right)$

Largest Eigenvalue of Wishart Matrix

- Let $X = [x_{i,j}] \rightarrow (M \times N)$ rectangular matrix
- $W = X^\dagger X \rightarrow (N \times N)$ square **covariance** matrix (**Wishart**)
- Entries of X Gaussian: $\text{Pr}[X] \propto \exp\left[-\frac{\beta}{2}\text{Tr}(X^\dagger X)\right]$
- N real eigenvalues of W : $\lambda_1 \geq 0, \lambda_2 \geq 0, \dots, \lambda_N \geq 0$
- Joint distribution of eigenvalues (James, 1964):

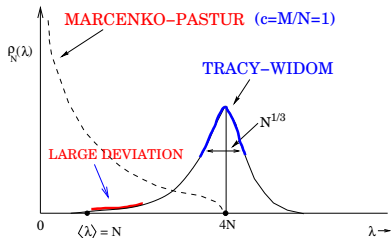
$$P(\{\lambda_i\}) \propto \exp\left[-\frac{\beta}{2}\sum_{i=1}^N \lambda_i\right] \prod_i \lambda_i^{\frac{\beta}{2}(1+M-N)-1} \prod_{j < k} |\lambda_j - \lambda_k|^\beta$$

- Av. density of states: $\rho(\lambda, N) = \left\langle \frac{1}{N} \sum_{i=1}^N \delta(\lambda - \lambda_i) \right\rangle \xrightarrow{N \rightarrow \infty} \frac{1}{N} f_{\text{MP}}\left(\frac{\lambda}{N}\right)$

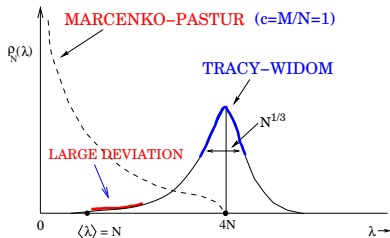
- Marcenko-Pastur law (1967): $f_{\text{MP}}(x) = \frac{1}{2\pi x} \sqrt{(x_+ - x)(x - x_-)}$

$$x_{\pm} = \left(1 \pm \frac{1}{\sqrt{c}}\right)^2 \text{ where } c = N/M \leq 1$$

Largest Eigenvalue of the Wishart Matrix

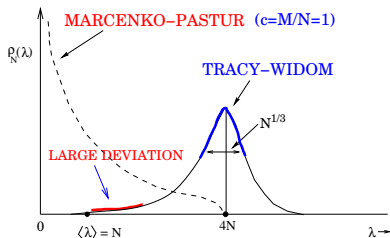


Largest Eigenvalue of the Wishart Matrix



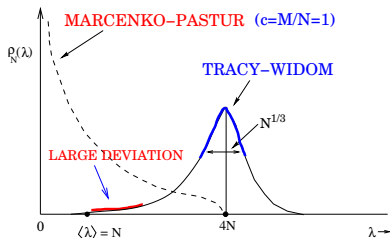
- average eigenvalue $\langle \lambda \rangle = N$ (e.g., for $M = N$)

Largest Eigenvalue of the Wishart Matrix



- average eigenvalue $\langle \lambda \rangle = N$ (e.g., for $M = N$)
- largest eigenvalue (e.g., for $M = N$): $\lambda_{\max} = 4N + 2^{4/3} N^{1/3} \Xi_{\text{TW}}$
valid for $|4N - \lambda_{\max}| \sim O(N^{1/3})$ (Johansson 2000, Johnstone 2001).

Largest Eigenvalue of the Wishart Matrix



- average eigenvalue $\langle \lambda \rangle = N$ (e.g., for $M = N$)
- largest eigenvalue (e.g., for $M = N$): $\lambda_{\max} = 4N + 2^{4/3} N^{1/3} \Xi_{\text{TW}}$
valid for $|4N - \lambda_{\max}| \sim O(N^{1/3})$ (Johansson 2000, Johnstone 2001).
- for $4N - t \sim O(N)$ (large negative fluctuation):

$$\text{Prob}[\lambda_{\max} \leq t, N] \sim \exp \left[-\beta N^2 \Phi_W \left(\frac{4N - t}{N} \right) \right]$$

Large Deviation for Wishart Matrix:

- large deviation function $\Phi_W(y) \rightarrow$ computed explicitly for all $c = N/M \leq 1$

(Vivo, S.M., & Bohigas, 2007)

for the special case $c = 1$ Deift et. al. computed $\Phi_W(y)$ by a different method (Deift, Its & Krasovsky, 2006).

Large Deviation for Wishart Matrix:

- large deviation function $\Phi_W(y) \rightarrow$ computed explicitly for all $c = N/M \leq 1$

(Vivo, S.M., & Bohigas, 2007)

for the special case $c = 1$ Deift et. al. computed $\Phi_W(y)$ by a different method (Deift, Its & Krasovsky, 2006).

- In particular, $P_N = \text{Prob}[\lambda_{\max} \leq \langle \lambda \rangle = N/c] \sim \exp[-\beta\theta(c)N^2]$

Large Deviation for Wishart Matrix:

- large deviation function $\Phi_W(y) \rightarrow$ computed explicitly for all $c = N/M \leq 1$

(Vivo, S.M., & Bohigas, 2007)

for the special case $c = 1$ Deift et. al. computed $\Phi_W(y)$ by a different method (Deift, Its & Krasovsky, 2006).

- In particular, $P_N = \text{Prob}[\lambda_{\max} \leq \langle \lambda \rangle = N/c] \sim \exp[-\beta\theta(c)N^2]$
- For example, for the square matrix case $M = N$, i.e., $c = 1$

Large Deviation for Wishart Matrix:

- large deviation function $\Phi_W(y) \rightarrow$ computed explicitly for all $c = N/M \leq 1$

(Vivo, S.M., & Bohigas, 2007)

for the special case $c = 1$ Deift et. al. computed $\Phi_W(y)$ by a different method (Deift, Its & Krasovsky, 2006).

- In particular, $P_N = \text{Prob}[\lambda_{\max} \leq \langle \lambda \rangle = N/c] \sim \exp[-\beta\theta(c)N^2]$
- For example, for the square matrix case $M = N$, i.e., $c = 1$

$$\theta(1) = \log(2) - \frac{33}{64} = 0.177522\dots$$

(Vivo, S.M., & Bohigas, 2007)

Large Deviation for Wishart Matrix:

- large deviation function $\Phi_W(y) \rightarrow$ computed explicitly for all $c = N/M \leq 1$

(Vivo, S.M., & Bohigas, 2007)

for the special case $c = 1$ Deift et. al. computed $\Phi_W(y)$ by a different method (Deift, Its & Krasovsky, 2006).

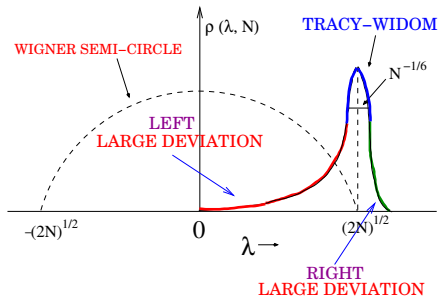
- In particular, $P_N = \text{Prob}[\lambda_{\max} \leq \langle \lambda \rangle = N/c] \sim \exp[-\beta\theta(c)N^2]$
- For example, for the square matrix case $M = N$, i.e., $c = 1$

$$\theta(1) = \log(2) - \frac{33}{64} = 0.177522\dots$$

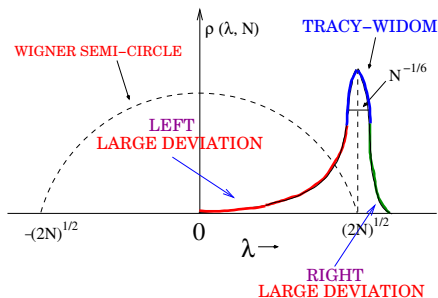
(Vivo, S.M., & Bohigas, 2007)

- Similarly, the **right** large deviation function (S.M. & Vergassola, 2009).

Summary and Conclusions

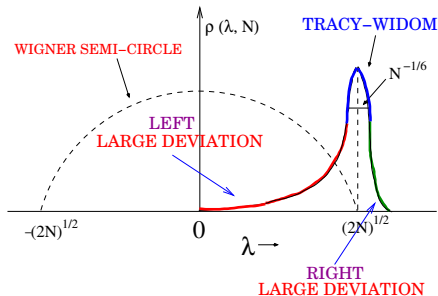


Summary and Conclusions



- Prob. distr. of $\lambda_{\max} \iff$ Coulomb gas with a wall

Summary and Conclusions

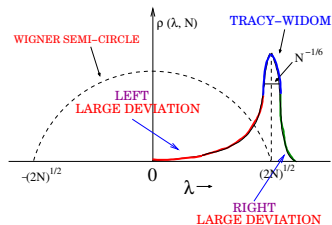


- Prob. distr. of $\lambda_{\max} \iff$ Coulomb gas with a wall
- for $(\sqrt{2N} - t) \sim N^{-1/6}$

$$\text{Prob}[\lambda_{\max} \leq t, N] \approx F_{\text{TW}} \left[N^{1/6} (t - \sqrt{2N}) \right]$$

Tracy-Widom Distribution

Summary and Conclusions: Large Deviations



Large Deviations:

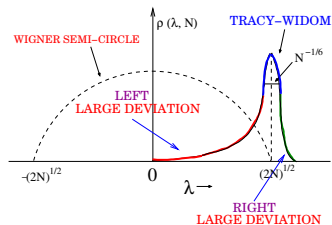
- for $(\sqrt{2N} - t) \sim \sqrt{N}$

$$\text{Prob}[\lambda_{\max} = t, N] \approx \exp \left[-\beta N^2 \Phi_- \left(\frac{\sqrt{2N} - t}{\sqrt{N}} \right) \right]$$

- for $(t - \sqrt{2N}) \sim \sqrt{N}$

$$\text{Prob}[\lambda_{\max} = t, N] \sim \exp \left[-\beta N \Phi_+ \left(\frac{t - \sqrt{2N}}{\sqrt{N}} \right) \right]$$

Summary and Conclusions: Large Deviations



Large Deviations:

- for $(\sqrt{2N} - t) \sim \sqrt{N}$

$$\text{Prob}[\lambda_{\max} = t, N] \approx \exp \left[-\beta N^2 \Phi_- \left(\frac{\sqrt{2N} - t}{\sqrt{N}} \right) \right]$$

- for $(t - \sqrt{2N}) \sim \sqrt{N}$

$$\text{Prob}[\lambda_{\max} = t, N] \sim \exp \left[-\beta N \Phi_+ \left(\frac{t - \sqrt{2N}}{\sqrt{N}} \right) \right]$$

Exact functional forms of $\Phi_-(y)$ (Dean & S.M., 2006) and $\Phi_+(y)$ (S.M. & Vergassola, 2009).

Conclusions and Related Problems

- In particular, $P_N = \text{Prob}[\lambda_1 \leq 0, \lambda_2 \leq 0, \dots, \lambda_N \leq 0] \sim \exp[-\beta\theta N^2]$

$$\theta = \frac{1}{4} \ln(3) = 0.274653.. \quad (\text{Dean \& S.M., 2006})$$

Conclusions and Related Problems

- In particular, $P_N = \text{Prob}[\lambda_1 \leq 0, \lambda_2 \leq 0, \dots, \lambda_N \leq 0] \sim \exp[-\beta\theta N^2]$

$$\theta = \frac{1}{4} \ln(3) = 0.274653.. \quad (\text{Dean \& S.M., 2006})$$

- Joint distribution of λ_{\min} and λ_{\max} (Dean & S.M., 2008)

Conclusions and Related Problems

- In particular, $P_N = \text{Prob}[\lambda_1 \leq 0, \lambda_2 \leq 0, \dots, \lambda_N \leq 0] \sim \exp[-\beta\theta N^2]$

$$\theta = \frac{1}{4} \ln(3) = 0.274653.. \quad (\text{Dean \& S.M., 2006})$$

- Joint distribution of λ_{\min} and λ_{\max} (Dean & S.M., 2008)
- Extension to **Wishart** matrices (Vivo, S.M. & Bohigas 2007, S.M. & Vergassola 2009), **Non-intersecting Brownian Interfaces near a wall** (C. Nadal & S.M., 2009)

Conclusions and Related Problems

- In particular, $P_N = \text{Prob}[\lambda_1 \leq 0, \lambda_2 \leq 0, \dots, \lambda_N \leq 0] \sim \exp[-\beta\theta N^2]$

$$\theta = \frac{1}{4} \ln(3) = 0.274653.. \quad (\text{Dean \& S.M., 2006})$$

- Joint distribution of λ_{\min} and λ_{\max} (Dean & S.M., 2008)
- Extension to **Wishart** matrices (Vivo, S.M. & Bohigas 2007, S.M. & Vergassola 2009), **Non-intersecting Brownian Interfaces near a wall** (C. Nadal & S.M., 2009)
- Technique \Rightarrow useful for other problems: Expected number of stationary points for Gaussian random fields (Bray & Dean 2007, Fyodorov et. al. 2007)

Conclusions and Related Problems

- In particular, $P_N = \text{Prob}[\lambda_1 \leq 0, \lambda_2 \leq 0, \dots, \lambda_N \leq 0] \sim \exp[-\beta\theta N^2]$

$$\theta = \frac{1}{4} \ln(3) = 0.274653.. \quad (\text{Dean \& S.M., 2006})$$

- Joint distribution of λ_{\min} and λ_{\max} (Dean & S.M., 2008)
- Extension to **Wishart** matrices (Vivo, S.M. & Bohigas 2007, S.M. & Vergassola 2009), **Non-intersecting Brownian Interfaces near a wall** (C. Nadal & S.M., 2009)
- Technique \Rightarrow useful for other problems: Expected number of stationary points for Gaussian random fields (Bray & Dean 2007, Fyodorov et. al. 2007)
- Large deviation of the conductance and shot noise in quantum dots (Vivo, S .M., & Bohigas 2008)

Conclusions and Related Problems

- In particular, $P_N = \text{Prob}[\lambda_1 \leq 0, \lambda_2 \leq 0, \dots, \lambda_N \leq 0] \sim \exp[-\beta\theta N^2]$

$$\theta = \frac{1}{4} \ln(3) = 0.274653.. \quad (\text{Dean \& S.M., 2006})$$

- Joint distribution of λ_{\min} and λ_{\max} (Dean & S.M., 2008)
- Extension to **Wishart** matrices (Vivo, S.M. & Bohigas 2007, S.M. & Vergassola 2009), **Non-intersecting Brownian Interfaces near a wall** (C. Nadal & S.M., 2009)
- Technique \Rightarrow useful for other problems: Expected number of stationary points for Gaussian random fields (Bray & Dean 2007, Fyodorov et. al. 2007)
- Large deviation of the conductance and shot noise in quantum dots (Vivo, S .M., & Bohigas 2008)
- Exact smallest eigenvalue distribution of an **entangled** random pure state (S.M., Bohigas, & Lakshminarayan 2008)

Collaborators

- O. Bohigas (Orsay, FRANCE)
- D.S. Dean (Toulouse, FRANCE)
- A. Lakshminarayan (IIT Chennai, INDIA)
- C. Nadal (Ph.D student, Orsay, FRANCE)
- M. Vergassola (Institut Pasteur, Paris, FRANCE)
- P. Vivo (ICTP, Trieste, ITALY)

References

- D.S. Dean and S.N. Majumdar, “ Large Deviations of Extreme Eigenvalues of Random Matrices”, *Phys. Rev. Lett.*, **97**, 160201 (2006)
- D.S. Dean and S.N. Majumdar, “Extreme Value Statistics of Eigenvalues of Gaussian Random Matrices”, *Phys. Rev. E.*, **77**, 041108 (2008)
- P. Vivo, S.N. Majumdar, and O. Bohigas, “Large Deviations of the Maximum Eigenvalue in Wishart Random Matrices”, *J. Phys. A. Math-Gen*, **40**, 4317 (2007)
- S.N. Majumdar, O. Bohigas and A. Lakshminarayan, “Exact Minimum Eigenvalue Distribution of an Entangled Random Pure State”, *J. Stat. Phys.*, **131**, 33 (2008)
- P. Vivo, S.N. Majumdar, and O. Bohigas, “Distributions of Conductance and Shot Noise and associated Phase Transitions”, *Phys. Rev. Lett.* **101**, 216809 (2008)
- S.N. Majumdar and M. Vergassola, “Large Deviations of the Maximum Eigenvalue for Wishart and Gaussian Random Matrices”, *Phys. Rev. Lett.* **102**, 060601 (2009)
- S.N. Majumdar, Review on Tracy-Widom and related problems, *Les Houches* lecture notes (2006)