

Assignment 5, due October 29

Corrections: (none yet.)

1. (A generalization of the Ito isometry formula) Here is a handy to calculate some things about Ito integrals

- (a) Suppose f_t and g_t are non-anticipating functions, and the corresponding Ito integrals are

$$X_t = \int_0^t f_s dW_s$$
$$Y_t = \int_0^t g_s dW_s .$$

Show that

$$\text{cov}(X_t, Y_t) = E[X_t Y_t] = \int_0^t E[f_s g_s] ds .$$

- (b) Suppose $f_t = t^2$ and $g_t = 1$. The notes for Week 5 show that $X_t \sim \mathcal{N}(0, t^5/5)$. Clearly $Y_t = W_t$. Compute the covariance of X_t and W_t using the result of part (a). This should agree with the result of question (3) from Assignment 3.
- (c) Since (X_t, W_t) is a bivariate normal whose variance/covariance structure you know, you can compute the conditional variance $\text{var}(X_t|W_t)$. Use this result to show that X_t is not a function of W_t . This is an example of a general phenomenon, that the value of an Ito integral depends on the whole path $W_{[0,t]}$, not just the endpoint W_t .
2. (Ornstein Uhlenbeck) This exercise goes through another approach to the Ornstein Uhlenbeck process. This time the process is called V_t because it represents the velocity of a small particle in a fluid at time t . This particle is subject to a random force F_t and friction with coefficient γ . We assume the units have been chosen so the mass of the particle is 1. The dynamics are

$$\frac{dV_t}{dt} = -\gamma V_t + F_t . \tag{1}$$

The term $-\gamma V_t$ represents friction proportional to the velocity of the particle, but pushing in the direction the particle is not moving. We always assume $\gamma > 0$.

- (a) Write the solution that corresponds to $F \equiv 0$ and $V_s = 1$. Call this $G(s, t)$. This plays the role that is played by a Green's function for a PDE.
- (b) Suppose F_t is a bounded function of t or grown slowly with t as $t \rightarrow -\infty$. Suppose that V_t likewise grows slowly with t or is bounded. Show that

$$V_t = \int_{-\infty}^t G(s, t) F_s ds . \quad (2)$$

Hint: Differentiate with respect to t . There are two terms, which correspond to the two terms on the right of (1).

- (c) An *impulsive* force of size 1 has the form $F_t = \delta(t - t_0)$. Describe the solution (2) with an impulsive force, both for $t < t_0$ and $t > t_0$.
- (d) Show that the formula (2) corresponds to a superposition of impulsive forces at time s over the interval ds with size $F_s ds$.
- (e) Suppose we replace the impulsive force over the interval $(s, s + ds)$ with a mean zero Gaussian σdW_s in (2). Find a formula for $E[V_t]$ and one for $\text{var}(V_t)$. Hint: These are independent of t , why? If you are uncomfortable with an infinite range of integration, you may replace $-\infty$ by a large negative t_0 in (2), then let $t_0 \rightarrow -\infty$.
- (f) Show that the V_t given by part (2e) is a Markov process. Hint: write a formula for $E[V_s | \mathcal{F}_t]$ that depends only on V_t and $W_{[t,s]}$.
- (g) Suppose $\Delta V = V_{t+\Delta t} - V_t$. Find a formula for $E[\Delta V | \mathcal{F}_t]$, and one for $E[(\Delta V)^2 | \mathcal{F}_t]$. Show that this is the same as the Ornstein-Uhlenbeck process in that the formulas here agree with the formulas from Week 3, Section 5 (possibly with 2γ for γ or 2σ for σ . Hint: for the latter, it may be simpler to calculate $\text{var}(\Delta V | \mathcal{F}_t)$, because the dependence on V_t is different.

3. In Einstein's model of Brownian motion, the location of a particle is

$$X_t = \int_0^t V_s ds . \quad (3)$$

This exercise shows that this is true, provided we use an appropriate scaling. The parameter γ from Exercise (2) controls how fast V_t loses memory. Therefore, in this exercise we take the limit $\gamma \rightarrow \infty$ and identify the limiting process X_t .

- (a) Find a formula for X_t in the form

$$X_t = \int_0^t L(s, t) dW_s .$$

Hint: Combine the two formulas (2) (with $F_s ds = dW_s$) and (3), reverse the order of integration.

- (b) Find a formula for σ as $\gamma \rightarrow \infty$ so that $E[X_1^2] = 1$. Call this process $X_{\gamma,t}$. Hint: the exact formula for finite γ may be hard to find, but you can find the behavior of σ as $\gamma \rightarrow \infty$ as a power of γ to leading order. This all you need.
- (c) Show that in the limit $\gamma \rightarrow \infty$ from part (3b) the process $X_{\gamma,t}$ has $X_{\gamma,[0,T]} \xrightarrow{D} X_t$ as $\gamma \rightarrow \infty$, where X_t is standard Brownian motion. Take this to mean that the finite dimensional joint distributions of $(X_{t_1}, \dots, X_{t_n})$ are what they should be for Brownian motion. Hint: Since $X_{\gamma,[0,T]}$ is Gaussian (being a linear function of $W_{[0,T]}$), you just need to evaluate the limiting means and covariances. You can do these from part (3b) and the independent increments property. So you need to show that as $\gamma \rightarrow \infty$, you approach independent increments.
- (d) (*only for those who have taken Probability Limit Theorems II or otherwise have the background to understand the question*) Complete part (3c) by showing that the $X_{\gamma,[0,T]}$ form a tight family. You can do this by finding uniform estimates of the form

$$E[\Delta X_\gamma^4] \leq C\Delta t^2,$$

which imply that the paths X_γ are uniformly Hölder continuous.

4. (*strong law of large numbers*) Suppose Y_k is a family of i.i.d. random variables with $E[Y_k] = \mu$. The *Kolmogorov strong law* of large numbers is the theorem that if $E[|Y_k|] < \infty$, then

$$\bar{Y}_n = \frac{1}{n} \sum_{k=1}^n Y_k \rightarrow \mu \text{ as } n \rightarrow \infty \text{ a.s.} \quad (4)$$

This exercise does not suggest his brilliant proof using the *three series lemma* or the more recent proof using the *Birkoff ergodic theorem*. Instead: Give a proof of (4) using the hypothesis $E[Y_k^4] < \infty$. Hint: Suppose $\mu = 0$. The statement $\bar{Y}_n \rightarrow 0$ is the same as the statement $\bar{Y}_n^4 \rightarrow 0$. Set $X_n = \bar{Y}_n^4$ and try to show that $\sum E[X_n] < \infty$ and use the Borel Cantelli style lemma from the notes. What do you do if $\mu \neq 0$?

5. (*Poisson process*) A simple *Poisson arrival process* is a sequence of times $0 = T_0 < T_1 < T_2 < \dots$. The *inter-arrival* times $S_k = T_k - T_{k-1}$ are independent exponential random variables. The *intensity* parameter, λ , is the parameter in the exponential distribution $S_k \sim \lambda e^{-\lambda s}$, $S_k > 0$. The *counting function*¹ is $N_t = k$ if $T_k < t$ and $T_{k+1} \geq t$. Either the counting

¹The inequality/equality choice makes the process N_t a *cadlag* process, more properly *càdlàg*, a French abbreviation of “continue à droite, limite à gauche”, which translates to “continuous on the right, limit on the left”. If you don’t know French, you can remember *droite*, which is related to the English word “right” (both as in “rights” and as a direction), and *gauche* is an English word related to being clumsy (or inappropriate), which is how it is for many of us with our left hand.

process or the arrival times are called the Poisson process. The counting process jumps from k to $k+1$ at time T_k . Therefore, it is sometimes called a *jump process*.

- (a) Derive the probability density, $f_k(t)$, of T_k . Hint:

$$P(T_k \in (t, t + dt)) = \int_{t'=0}^t f_{k-1}(t') P(T_k \in (t, t + dt) | T_{k-1} = t') dt' .$$

Write this in terms of the S_k density, Figure out the integrals, starting with $f_1(t) = \lambda e^{-\lambda t}$, then moving to $f_2(t)$, f_3 , etc., until you see the pattern.

- (b) Derive a formula for $p_n(t) = P(N_t = n)$. This is the *Poisson* distribution. Check that your formula satisfies $\sum_0^\infty p_n(t) = 1$. This involves the Taylor series formula for the exponential. Hint: $p_0(t) = P(T_1 > t)$, $p_1(t) = P(T_1 < t < T_2)$, etc. Look for the pattern. Prove it by induction.
- (c) Introduce a small time increment Δt and a probability $p_{\Delta t} = \lambda \Delta t$ ($p_{\Delta t} < 1$ for Δt small enough). Define $t_j = j \Delta t$ and independent random variables $Y_j = 1$ with probability $p_{\Delta t}$ and $Y_j = 0$ otherwise. Define

$$N_t^{\Delta t} = \sum_{t_j < t} Y_j .$$

Show that for each t ,

$$N_t^{\Delta t} \xrightarrow{\mathcal{D}} N_t \quad \text{as } \Delta t \rightarrow 0 .$$

Hint: The distribution of $N_t^{\Delta t}$ is binomial. The limit $\Delta t \rightarrow 0$ is easy for $p_n^{\Delta t}(t)$.

- (d) Assuming that $N_{[0,T]}^{\Delta t} \xrightarrow{\mathcal{D}} N_{[0,T]}$ as $\Delta t \rightarrow 0$, show that the Poisson process has the independent increments property. Hint: this is a statement about discrete probabilities that you can check by using those probabilities, and figuring out what happens if t is not one of the t_j .
- (e) Show that N_t is a Markov process.
- (f) The *compensated* Poisson arrival process is $M_t = N_t - t$. Show that this is a martingale, which means that if $s > t$, then

$$E[M_s | \mathcal{F}_t] = M_t .$$

- (g) The standard Poisson process has intensity, or *arrival rate*, $\lambda = 1$. Show that

$$E[\Delta M^2 | \mathcal{F}_t] = \Delta t + (\text{smaller}) \quad \text{as } \Delta t \rightarrow 0 .$$

As usual, $\Delta M = M_{t+\Delta t} - M_t$. Compare these to comparable facts about standard Brownian motion (martingale, $E[\Delta W^2 | \mathcal{F}_t]$). Conclude that Brownian motion is not the unique process with these properties.

- (h) Calculate the scaling of $E[\Delta W^4 | \mathcal{F}_t]$ and $E[\Delta M^4 | \mathcal{F}_t]$ with Δt as $\Delta t \rightarrow 0$.